

### Design Alternatives for Integrating the National Medical Expenditure Survey With the National Health Interview Survey

Research was undertaken to evaluate alternative methods of selecting a sample of eligible respondents for the National Medical Expenditure Survey (NMES) from the National Health Interview Survey (NHIS). This report presents estimates of the effects of alternative design options, obtained by statistical modeling techniques, for linking the NMES with the NHIS. The estimated survey costs for alternative linked and unlinked design options are compared for fixed precision. The findings indicate that substantial savings would be realized by linking the NMES to the NHIS if a premium is put on small-domain estimates.

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### Foreword

This is the second report presenting results of research on the effects of integrating the designs of the National Center for Health Statistics (NCHS) national household sample surveys, which heretofore were designed as independent surveys. Design integration would be accomplished by using the files of the National Health Interview Survey (NHIS), the largest and only continuing NCHS population survey, as the sampling frame for NCHS's other population surveys. Research findings with respect to linking the 1987 National Survey of Family Growth (NSFG) to NHIS were presented in an earlier report in this publication series, and the findings relating to the 1987 National Medical Expenditure Survey (NMES) are presented in this report.

The earlier report indicated that significant economies would be realized by linking NSFG to NHIS because NSFG requires a substantial oversampling of households with black females. However, it was unreasonable to assume that the NSFG findings would necessarily apply to NMES because NSFG is a single-time retrospective survey and NMES is a panel survey. As such, the population domains of interest would be different for NMES and NSFG. As it turned out, the NMES and NSFG research findings were quite similar. Among other things, this report concludes that substantial savings would be realized by linking NMES to NHIS if NMES puts a premium on small-domain estimates.

I provided technical oversight to this project, which was conducted under a contract with the Research Triangle Institute. Dr. Andrew White was instrumental in guiding this report through the publication process by working closely with the authors and the editors.

> Monroe G. Sirken Associate Director for Research and Methodology

### Symbols

- --- Data not available
- ... Category not applicable
- Quantity zero
- 0.0 Quantity more than zero but less than 0.05
- Quantity more than zero but less than
   500 where numbers are rounded to thousands
- \* Figure does not meet standard of reliability or precision
- # Figure suppressed to comply with confidentiality requirements

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# Design Alternatives for Integrating the National Medical Expenditure Survey With the National Health Interview Survey

by Brenda G. Cox, Ralph E. Folsom, and Thomas G. Virag, Research Triangle Institute

## Chapter 1 Introduction

Current planning for population-based surveys conducted by the National Center for Health Statistics (NCHS) suggests that the data systems can be integrated to save on data collection costs, to reduce respondent burden, and to increase the utility of the resultant data. As part of the NCHS effort to evaluate advantages of an integrated data system, Research Triangle Institute examined alternative designs for integrating the National Medical Expenditure Survey (NMES) with the larger National Health Interview Survey (NHIS). NMES will be a longitudinal study of the 1987 health care utilization and expenditures of civilian noninstitutionalized residents of the United States. This report summarizes the results of an investigation to assess the feasibility of linking the two surveys.

As a baseline for comparison, specifications for an unlinked NMES design were developed. Selected independently of NHIS, this unlinked design results in a stratified, clustered area sample similar to that of the 1980 National Medical Care Utilization and Expenditure Survey. For flexibility of NCHS planning, two sample sizes were used: 6,000 and 10,000 responding households. The 6,000-household design is similar in size to the 1980 National Medical Care Utilization and Expenditure Survey. The 10,000-household design was added so that NCHS could evaluate the improved precision for surveying smaller domains with the larger sample against the increased survey cost. Survey costs for the two sample size alternatives were modeled as well as the variances for selected statistics of interest.

The second design for which specifications were developed was a linked dwelling unit design. The linked dwelling unit design selects the sample of individuals to be included in NMES by subsampling NHIS sample dwelling units. In round 1 of NMES, the occupants of the subsampled dwelling units would be interviewed. Rounds 2–5 of date collection would use the same procedures as the unlinked NMES design. To measure the effect of the number of NHIS primary sampling units (PSU's) from which the NMES sample dwelling units are selected, both a 100-PSU and a 200-PSU linked dwelling unit design were investigated. For each design, two sample size alternatives were also investigated. These two sample sizes are those required to yield the same precision as the unlinked design with 6,000 and 10,000 responding households. The third set of specifications developed were for a linked household design. The linked household design selects a sample of NHIS households for inclusion in NMES. The individuals within the subsampled households are interviewed in round 1 whether or not they live in the clustered NHIS sample dwelling units. Rounds 2-5 data collection uses the same rules as the unlinked design. As in the linked dwelling unit design, to assess the effect of the number of PSU's, designs were developed for both 100 PSU's and 200 PSU's; two sample sizes were investigated. These sample sizes were determined as the sizes required to yield the same precision as the unlinked design with 6,000 and 10,000 responding households.

Each of these designs is self-weighting; that is, all sample individuals are selected with the same probability. In many ways this eliminates the chief advantage of linkage with NHIS. With knowledge of individual characteristics available for NHIS sample respondents, added precision can be obtained for small domains without proportionally increasing the size of the total sample. To evaluate this feature of NHIS linkage, a fourth and final design type was investigated. This design is an optimally allocated linked household design in which the precision constraints set for the total population and the Medicaid population were based on those achieved by the unlinked design. Instead of arbitrarily determining the number of NHIS PSU's and segments to include, optimal sizes were determined for these components.

The development of these four designs is described in the following chapters. An important finding of this investigation is that there appears to be little relative gain from linkage when the final design is self-weighting. The principal gain from the linked self-weighting design is in the elimination of costs associated with counting and listing. Because the NMCUES interview pattern for all rounds was adopted in this investigation (personal interviews are used in the first two rounds and telephone interviews in the third and fourth rounds), there is little gained from the names, addresses, and telephone numbers of NHIS sample individuals. The optimally allocated design, however, uses characteristics of NHIS respondents to oversample heavy users of health care services and to increase the precision for small domains without proportionally increasing the size of the total sample.

# Chapter 2 The unlinked National Medical Expenditure Survey design

The unlinked National Medical Expenditure Survey (NMES) designs studied in this investigation were patterned after the design used for the 1980 National Medical Care Utilization and Expenditure Survey (NMCUES). Specifically, an area sampling approach was used incorporating a selfweighting design in which each sample individual is selected with equal probability. The sample sizes required to yield 6,000 and 10,000 responding households were determined as well as the survey costs associated with these designs. The variances achieved by the unweighted, unlinked NMES design were modeled for use in sample size determination for the remaining designs.

### Definition

The unlinked sample design is a stratified, multistage area probability design in which each sample dwelling unit is selected with equal probability. (In this report, the term "dwelling unit" refers to either a housing unit or a group quarters listing unit.) The first-stage sample consists of primary sampling units (PSU's) that are counties, parts of counties, or groups of contiguous counties. The second-stage sample consists of secondary sampling units that are census enumeration districts or block groups. Smaller area segments constitute the third stage. All of the dwelling units within these sample segments are listed. During the fourth stage of sampling, dwelling units within these sample segments are designated for inclusion in the NMES sample.

All civilian noninstitutionalized individuals residing in the sampled dwelling units in round 1 are included in the survey. Single college students in the 17–22-year age range are linked to their parents' residence and included in the survey only when their parents' residence is selected. Round 1 data collection uses personal interviews except for college students living outside a 2-hour, one-way drive of a sample PSU. In this case, telephone interviewing is used.

In round 2, these key persons are interviewed in their round 2 location. Individuals and families that moved must be traced to determine their new addresses. Individuals who joined the family of a key individual by birth or return from an institution, the military, or an overseas residence are included in NMES as a key person. Other individuals joining the families of key persons are classified as nonkey. Data are collected for both key and nonkey persons. The data for key persons are needed for person-level analyses. The data for nonkey persons are needed for family-level analyses only. Data collection in round 2 also uses personal interviews except for college students and movers outside a 2-hour, one-way drive from a sample PSU.

In round 3, data collection is primarily by telephone, with personal interviews conducted only for households without telephones and households requesting personal interviews. Key persons who move from their round 2 locations must be traced and interviewed at their new locations. Nonkey persons who moved are interviewed only when a key person moves with them. Individuals who are born or who return from an institution, the military, or overseas residence are included as key persons. Other individuals joining the families of key persons are classified as nonkey; data are gathered for them only during the time in which they were members of a key person's family.

The mode of data collection in round 4 follows that of round 3 with similar guidelines for key and nonkey persons. Because December 31 is the end of the survey reference period, approximately 30 percent of the sample is not interviewed in round 4 but instead early in round 5 (that is, shortly after January 1 of the next year).

The final round of data collection primarily uses personal interviewing under the same guidelines used in previous rounds to define key and nonkey persons and to determine movers who will be followed.

### Sample size determination

Two sets of sample sizes were required for the unlinked NMES design: A sample size sufficient to yield 6,000 responding households, and a sample size sufficient to yield 10,000 responding households. To obtain these sizes, a precise definition was needed for "responding household." It was decided to use responding originating base reporting units (OBRU's) and to describe the sample sizes needed as those yielding an OBRU design with 6,000 responding and an OBRU design with 10,000 responding. These OBRU's are the round 1 reporting units (RU's) after college student RU's are linked back to parent RU's. Because data collection costs relate to reporting units (RU's) and rounds, sample sizes in terms of these units were developed.

The first step in this process was to model the 1980 NMCUES experience starting with the set of control system records generated by responding OBRU's. (In the NMCUES, an OBRU was defined to be responding if it was linked to an RU that completed an interview in any of the five data collection rounds.) The NMCUES contained 6,269 responding OBRU's. These responding OBRU's generated 6,603 completed RU interviews in round 1, 6,519 completed RU interviews in round 2, 6,528 completed RU interviews in round 3, 4,559 completed RU interviews in round 4, and 6,561 completed RU interviews in round 5. These were more RU interviews than there were responding OBRU's because OBRU's containing college students required more than one RU assignment to handle the different addresses at which data collection occurred. The NMCUES interviews occurred in 135 PSU's and 809 segments.

Because the NMES should experience no worse than the nonresponse and attrition encountered by the 1980 NMCUES, the NMCUES experience was ratio adjusted to produce the sample sizes required for the OBRU designs with 6,000 and 10,000 responding. These sample sizes are summarized in table 1. For modeling convenience, it was assumed that the Research Triangle Institute (RTI) General Purpose Sample would be used, which contains 102 PSU's. The average segment size was set to the 1980 NMCUES experience of eight responding OBRU's. With eight responding OBRU's per segment, the OBRU design with 6,000 responding would require 750 segments, and the OBRU design with 10,000 responding would require 1,250 segments.

#### Variance modeling

As a baseline for comparison of the unlinked with the linked designs, the precision of the linked designs was fixed to that of the unlinked design for selected key statistics and key domains. The designs were then compared with respect to sample sizes and costs. The domains of interest were the total population, those individuals below 150 percent of poverty, Medicare recipients, Medicaid recipients, and individuals from families with college-educated heads of households. The statistics of interest were as follows:

- Average number of hospital visits.
- Average number of facility visits.
- Average number of office visits.
- Average annual expenditure for hospital visits.
- Average annual expenditure for facility visits.
- Average annual expenditure for office visits.
- Average annual out-of-pocket expense for hospital visits.
- Average annual out-of-pocket expense for facility visits.
- Average annual out-of-pocket expense for office visits.
- Proportion with large out-of-pocket expenditures.

To determine the sample sizes required for the linked designs, the variance was modeled for the OBRU unlinked, self-weighting designs with 6,000 and 10,000 responding using the 1980 NMCUES data.

The NMES estimation approach constructs means in terms of total person-years rather than in terms of all persons ever existing in the data collection year. For domain k, the mean utilization or expenditure per person-year is estimated as

$$\overline{Y}_{k}(\text{NMES}) = \frac{\sum_{i \in S} W(i)\delta_{k}(i)Y(i)}{\sum_{i \in S} W(i)T(i)\delta_{k}(i)}$$
(1)

where W(i) = analysis weight for the *i*th person

- $\delta_k(i) = 1$  if the *i*th person belongs to the *k*th domain and 0 if not
- Y(i) = response of the *i*th person
- T(i) = time-adjustment factor for the *i*th person

The numerator estimates total expenditures or utilization and the denominator the average annual number of persons in the population (that is, the total person-years). The time-adjustment factor T(i) is the total days that person *i* is eligible divided by the number of days in the year.

Large out-of-pocket expenditures are defined as "annualized" out-of-pocket expenditures of \$200 or more. The annualized out-of-pocket expenditure is the annual out-of-pocket expenditure divided by the fraction of the year during which the person is eligible. For domain k, the proportion with large outof-pocket expenditures is estimated as

$$\overline{Y}_{k}(\text{NMES}) = \frac{\sum_{i \in S} W(i)T(i)\delta_{k}(i)Y(i)}{\sum_{i \in S} W(i)T(i)\delta_{k}(i)}$$
(2)

where Y(i) = 1 if the person had large out-of-pocket expenditures and 0 if not.

The variables used in constructing these estimates were interim variables from the NMCUES analysis files and not the final variables contained in the public use files. For this reason, the estimates in this report may differ from those in other NMCUES reports.

The variance of  $\overline{Y_k}$  (NMES) was derived assuming a threestage household survey design patterned after the 1980 NMCUES sample design with PSU's of standard metropolitan statistical area, or county-size and area segments (SEG's) selected as noncompact clusters of dwelling units. The households containing at least one RU response are designated as responding OBRU's. Using this approach, the variance of  $\overline{Y_k}$  (NMES) may be modeled as

$$\operatorname{Var}\left[\overline{Y}_{k}(\operatorname{NMES})\right] = \frac{\sigma_{k}^{2}(\operatorname{PSU})}{r} + \frac{\sigma_{k}^{2}(\operatorname{SEG})}{r\overline{s}} + \frac{\sigma_{k}^{2}(\operatorname{OBRU})}{r\overline{s}\overline{t}} \quad (3)$$

where  $\sigma_k^2(\text{PSU}) =$  between-PSU, within-stratum variance component for domain k

r = number of PSU's

- $\sigma_k^2(SEG) =$  between-segment, within-PSU variance component for domain k
  - $\overline{s}$  = average number of segments per PSU
- $\sigma_k^2(\text{OBRU}) = \text{between-OBRU}$ , within-segment variance component for domain k
  - $\bar{t}$  = average number of responding OBRU's per segment

The variance components were estimated using 1980 NMCUES data.

The variance components estimation program, developed at RTI by Shah<sup>1</sup> for evaluating the efficiency of complex sample designs, was applied to the NMCUES data to produce the generalized composite components for PSU's, segments (SEG's), and OBRU's. VMCPNLS estimates the composite variance components in terms of an expression for the variance of a multistage Horvitz-Thompson estimator derived by Gray.<sup>2</sup> For the NMCUES design, VMCPNLS yields a four-stage analysis including a between-PSU component [ $\sum_{k}^{2}$ (PSU)]; a between-segment, within-PSU component [ $\sum_{k}^{2}$ (OBRU)]; a between-OBRU, within-segment component [ $\sum_{k}^{2}$ (OBRU)]; and a between-person (PID), within-OBRU component [ $\sum_{k}^{2}$ (PID)].

Because there is no subsampling of household members in NMCUES, the four-stage decomposition produced by VMCPNLS must be converted to the three-stage decomposition specified in equation (3). With the four-stage model, the PSU and segment components are equivalent to the corresponding parameters of the three-stage model. The OBRU-level component can be estimated from the four-stage components as  $\sum_{k}^{2}(OBRU) + \sum_{k}^{2}(PID)/\overline{n}$  where  $\overline{n}$  is the average number of responding persons per responding OBRU. Using the 1980 NMCUES data,  $\overline{n}$  is estimated to be 2.73.

The variance components estimated using the 1980 NMCUES data contain an effect due to unequal weighting of the NMCUES sample. To remove the unequal weighting effect, these components were converted to the variance proportions  $\Delta_k$ (PSU),  $\Delta_k$ (SEG), and  $\Delta_k$ (OBRU) by dividing by the total variation or

$$\Delta_{k}(\text{PSU}) = \frac{\sum_{k=2}^{2} (\text{PSU})}{\sum_{k=2}^{2} (\text{TOT})}$$
(4)

$$\Delta_k(\text{SEG}) = \frac{\sum_{k=1}^{2} (\text{SEG})}{\sum_{k=1}^{2} (\text{TOT})}$$
(5)

$$\Delta_{k}(\text{OBRU}) = \frac{\sum_{k}^{2}(\text{OBRU}) + \sum_{k}^{2}(\text{PID})/\overline{n}}{\sum_{k}^{2}(\text{TOT})}$$
(6)

where  $\sum_{k=1}^{2} (TOT)$  is defined as

$$\sum_{k}^{2} (\text{TOT}) = \sum_{k}^{2} (\text{PSU}) + \sum_{k}^{2} (\text{SEG}) + \sum_{k}^{2} (\text{OBRU}) + \frac{\sum_{k}^{2} (\text{PID})}{\frac{\pi}{n}}$$
(7)

Table 2 displays these variance proportions for the 5 domains of interest and the 10 outcome measures described earlier.

To obtain the  $\sigma^2$  variance components used in modeling the variance of the key statistics, the variance proportions were multiplied by the estimated population variance for the kth domain, denoted by  $S^2(k)$ . That is,

$$\sigma_k^2(\text{PSU}) = \Delta_k(\text{PSU})S^2(k) \tag{8}$$

$$\sigma_k^2(\text{SEG}) = \Delta_k(\text{SEG})S^2(k) \tag{9}$$

$$\sigma_k^2(\text{OBRU}) = \Delta_k(\text{OBRU})S^2(k) \tag{10}$$

A Taylor series approximation for the simple random sampling variance of a combined ratio estimator was used to estimate  $S^2(k)$ . The numerator was the Y total for domain k and the denominator the total person-years for domain k. (See equations (1) and (2).)

These three-stage variance component estimates were used to estimate the variances that would be achieved by self-weighting NMES OBRU designs with 6,000 and 10,000 responding. The terms remaining to be specified in the variance expression presented in equation (3) are the number of PSU's, r; the average number of segments sampled per PSU,  $\bar{s}$ ; and the average number of OBRU's sampled per segment,  $\bar{t}$ . For modeling purposes, the RTI's General Purpose Sample was assumed, which contains 102 PSU's (r = 102). Because the 1980 NMCUES had been designed to be optimal with respect to the number of selections per segment, the number of responding OBRU's per segment was set to the value that the 1980 NMCUES achieved, or  $\bar{t} = 8$ . Therefore, the total number of segments in the OBRU design with 6,000 responding would be 750 ( $r\bar{s} = 750$ ) and 1,250 for the OBRU design with 10,000 responding ( $r\bar{s} = 1,250$ ).

These estimated variances were used as precision criteria for the other designs investigated in this study. Table 3 presents the results of this variance modeling activity for the 5 domains of interest and the 10 outcome measures. For convenience, percent relative standard errors are used rather than the variances. The percent relative standard error is 100 times the standard error (the square root of the variance) divided by the parameter being estimated. The percent relative standard errors achieved by the OBRU design with 6,000 responding are sufficient for the estimates based upon the total domain, but the increased precision that the OBRU design with 10,000 responding achieves for the small domain estimates is desirable.

#### Cost modeling

To establish cost comparisons between the unlinked and the linked designs, a systematic method was developed to generate the costs for all designs. The approach used was to develop unit costs by task for each design. The NMES tasks included in the modeling were the basic sampling and weighting tasks and the data collecting and processing tasks:

- Survey sampling.
- Instrument and materials development.
- Field preparations.
- Survey training.
- Data collection.
- Control system development and production.
- Data receipt, editing, and document control.
- Data coding operations.
- Data entry operations.
- Control card development, maintenance, and production.
- Summary development, maintenance, and production.
- Other data processing operations.
- Database construction.
- Counting and listing.
- Project administration.

The unit costs that were developed for each task were fixed costs, PSU-level costs, segment-level costs, and reporting-unit-level costs.

The first step in the process was to document the RTI cost experience for the 1980 NMCUES. Because of insufficient data for other contractors' costs, modeling was conducted with only RTI data. Only direct costs were included in the modeling because indirect costs, such as the costs for administration and building maintenance, vary among contractors as do accounting procedures used to recover these costs. Another step in documenting RTI costs for NMCUES was to separate the National Household Survey (HHS) costs from the costs associated with the four State Medicaid Household Surveys (SMHS). In most cases, SMHS activity was conducted under task numbers different from the HHS. In situations where HHS data and SMHS data were processed simultaneously, the additional costs added by SMHS were removed.

The next step was to use the 1980 NMCUES cost experience to develop unit costs for each task. Derivation of the unit costs by NMES task was a time-consuming process. The appendix includes a discussion of this process. The results are summarized in tables 4 and 5. Table 4 presents the costs for the OBRU design with 6,000 responding by category of cost for each of the 15 NMES tasks. Table 5 presents the costs for the OBRU design with 10,000 responding. For the OBRU design with 6,000 responding, direct costs are \$4,963,013. For the OBRU design with 10,000 responding, direct costs are \$7,209,409.

#### Other design considerations

Data for the 1980 NMCUES were collected by two contractors: RTI and the National Opinion Research Center (NORC). The cost modeling presented in this chapter was based on data from one contractor, however. There are advantages and disadvantages associated with using more than one contractor in data collection. These differences include quality, timeliness, and cost considerations.

Whether the OBRU design with 10,000 responding is chosen over the OBRU design with 6,000 responding, NMES will have time constraints on data collecting and processing, because data collection rounds are approximately 3 months apart. In the time between rounds 2 and 3, for instance, the data for round 2 must be collected, keyed, edited, coded, and entered into the database. The database is then used to generate a cumulative summary of household health care utilization and expenditures. This summary must be mailed to each household and interviewer before round 3. The volume of data collecting and processing required in this limited timeframe is beyond the capability of all but the largest firms. Hence, many firms would need to work together to accomplish the task.

Another advantage of using more than one contractor is the potential for improvements in work quality. Access to experienced interviewing and supervisory staff is limited to the volume of work performed. The inhouse staff needed to monitor data collection, to edit and to key the data, and to produce the final database is also limited. Merging the resources of more than one contractor enlarges the pool of experienced staff who can be assigned to a task.

The disadvantage of using more than one contractor is the inevitable duplication of effort. Each organization incurs the fixed costs associated with sampling, data collection, and data processing. To determine the cost penalty of using two contractors, the cost model that had been developed to determine costs for the 1980 NMCUES if only RTI had done the survey was used. The sample sizes of the 1980 NMCUES were used with one exception. Although the survey included 135 PSU's, only 108 were unique. Because overlapping of PSU's between the general purpose samples of the contractors was a duplication of effort, RTI-only 1980 NMCUES costs were modeled using 108 PSU's.

Table 6 summarizes the results of this comparison. RTI and NORC tasks were consolidated so that they correspond closely; therefore, the costs presented in this comparison are estimated costs. For example, many of the NORC tasks involved HHS and SMHS. Because the data collection instrument was the same for the surveys, both contractors combined the data entry and data processing tasks for HHS and SMHS. These tasks were adjusted by the number of the total that were HHS. RTI was responsible for the development of many procedures and materials used by both contractors. These development costs as well as the maintenance and production costs are contained in the RTI costs for the control system, control card, and summary. RTI keyed much of the data that NORC collected. Because this activity was performed under a separate charge number, the costs for RTI keying of NORC data are entered in the NORC column. Both contractors used their general purpose half-samples, so there were minimal costs for counting and listing. If RTI had done the full NMCUES, additional counting and listing would have been required for the portion of the RTI half-sample not in routine use. These costs have been included under the data collection task. Finally, database construction was performed exclusively by RTI and

printing by NORC, so these tasks are listed as separate entries with zero costs for the other contractor.

Examination of table 6 suggests that there is indeed a substantial cost penalty associated with the use of two contractors for NMCUES. This examination estimates the cost of using two contractors for the 1980 NMCUES as a \$1,157,658 increase in direct costs for the study or an 18-percent increase over the costs for one contractor. The primary reason for the cost increase is that both contractors must incur fixed costs for sampling, data collection, and data processing. However, the *capability* of a single contractor to achieve results equivalent to NMCUES must be considered in weighing the advantages and disadvantages of using one versus two contractors.

Table 1. Completed reporting unit interviews by round for the unlinked designs with 6,000- and 10,000-respondent originating base reporting units (OBRU's)

Round	1980 NMCUES	6,000 respondent OBRU's	10,000-respondent OBRU's
1	6,603	6,319	10,531
2	6,519	6,238	10,397
3	6,528	6,247	10,411
4	4,559	4,363	7,271
5	6,561	6,278	10,464

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# Table 2. Proportions of National Medical Care Utilization and Expenditure Survey (NMCUES) expenditures and utilization variation by domain and type of service

	Pro	Proportion of variation <sup>1</sup>		
Domain and outcome measure	Δ( <i>PSU</i> )	$\Delta(SEG)$	∆(obru	
Total				
/isits:				
Hospital	0.0061	0.0007	0.9932	
Facility	0.0134	0.0517	0.9349	
Office	0.0066	0.0202	0.9732	
Charges:				
Hospital	0.0002	0.0028	0.9970	
Facility	0.0059	0.0338	0.9603	
Office	0.0003	0.0328	0.9669	
xpenses:				
Hospital, out of pocket (OOP)	0.0002	0.0065	0.9933	
Facility, OOP	0.0048	0.0092	0.9860	
Office, OOP	0.0002	0.0631	0.9367	
roportion with large OOP expenses	0.0002	0.0593	0.9405	
150 percent of poverty population				
isits:				
Hospital	0.0002	0.0117	0.9881	
Facility	0.0002	0.0557	0.9441	
•				
Office	0.0038	0.0279	0.9683	
harges:	0.0002	0.0131	0.9867	
Hospital	0.0002		0.9867	
Facility	0.0003	0.0456		
Office	0.0052	0.0262	0.9686	
xpenses: Hospital, OOP	0.0002	0.0002	0.9996	
Facility, OOP	0.0002	0.0113	0.9885	
Office, OOP	0.0002	0.0277	0.9721	
Proportion with large OOP expenses	0.0002	0.0248	0.9750	
Medicare recipients				
/isits:				
Hospital	0.0039	0.0002	0.9959	
Facility	0.0003	0.0003	0.9994	
Office	0.0114	0.0003	0.9883	
Sharges:	0.0111	0.0000	0.0000	
Hospital	0.0035	0.0005	0.9960	
Facility	0.0003	0.0003	0.9994	
Office	0.0081	0.0033	0.9886	
xpenses:	0.0001	0.0000	0.0000	
Hospital, OOP	0.0002	0.0002	0.9996	
Facility, OOP	0.0008	0.0003	0.9989	
Office, OOP	0.0095	0.0198	0.9707	
Proportion with large OOP expenses	0.0002	0.0137	0.9861	
	0.0002	0.0107	0.0001	
Medicaid recipients				
isits:				
Hospital	0.0007	0.0073	0.9920	
Facility	0.0041	0.0360	0.9599	
Office	0.0049	0.0056	0.9898	
iharges:	0.0000	0.0000	o oo	
Hospital	0.0002	0.0083	0.9918	
Facility	0.0003	0.0153	0.9844	
Office ,	0.0050	0.0002	0.9948	
		0.0000	c cc	
xpenses:	0 0000	0.0003	0.9978	
xpenses: Hospital, OOP	0.0019			
xpenses: Hospital, OOP Facility, OOP	0.0003	0.0003	0.9994	
xpenses: Hospital, OOP			0.9994 0.9978 0.9769	

<sup>1</sup>PSU = primary sampling unit; SEG = area segment; OBRU = originating base reporting unit.

Table 2. Proportions of National Medical Care Utilization and Expenditure Survey (NMCUES) expenditures and utilization variation by domain and type of service---Con.

		Proportion of variation <sup>1</sup>		
Domain and outcome measure	Δ( <i>PSU</i> )	$\Delta(SEG)$	$\Delta(OBRU)$	
College head of household population				
Visits:				
Hospital	0.0017	0.0020	0.9963	
Facility	0.0056	0.0333	0.9611	
Office	0.0002	0.0155	0.9843	
Charges:				
Hospital	0.0008	0.0075	0.9917	
Facility	0.0053	0.0003	0.9944	
Office	0.0002	0.0175	0.9822	
Expenses:				
Hospital, OOP	0.0001	0.0119	0.9880	
Facility, OOP	0.0003	0.0266	0.9731	
Office, OOP	0.0002	0.0329	0.9669	
Proportion with large OOP expenses	0.0012	0.0150	0.9838	

<sup>1</sup>PSU = primary sampling unit; SEG = area segment; OBRU = originating base reporting unit.

Table 3. Esti	timated means and relative standard errors for the unlinked National Medical Expenditure Survey (NMES) design with 6,000- and
10,000-respor	ndent originating base reporting units (OBRU's)

		Relative standard error		
Domain and outcome measure	Y <sub>k</sub> (NMES)	6,000-respondent OBRU's	10,000-respondent OBRU's	
Total				
Visits:	0.10	0.14	0.01	
Hospital	0.18	3.11	2.61	
	0.86	4.92	4.25	
Office	4.18	2.02	1.69	
Charges:				
Hospital	362.04	6.22	4.84	
Facility	50.56	4.95	4.11	
Office	117.71	2.42	1.88	
Expenses:				
Hospital, out of pocket (OOP)	33.10	12.08	9.39	
Facility, OOP	9.77	4.82	3.99	
Office, OOP	53.70	2.43	1.89	
Proportion with large OOP expenses.	0.24	7.03	5.47	
	0.24	7.03	0.47	
150 percent of poverty population				
/isits:				
Hospítal	0.24	5.29	4.11	
Facility	1.22	8.33	6.47	
Office	4.23	4.10	3.34	
Charges:	4.20	4.10	0.54	
	E16 02	12.04	10.14	
Hospital	516.93	13.04	10.14	
Facility	66.65	10.87	8.45	
Office	108.82	4.79	3.95	
xpenses:				
Hospital, OOP	40.40	15.31	11.91	
Facility, OOP	9.70	8.50	6.61	
Office, OOP	38.82	5.46	4.24	
Proportion with large OOP expenses,	0.20	13.55	10.53	
Medicare recipients				
/isits:				
Hospital	0.40	5.74	4 72	
Facility		9.97		
	1.45		7.76	
Office	7.27	4.38	3.83	
Charges:				
	1,164.15	11.18	9.14	
Facility	88.14	12.81	9.97	
Office	212.31	7.17	6.11	
xpenses:				
Hospital, OOP	79.02	17.82	13.87	
Facility, OOP	13.47	10.47	8.23	
Office, OOP	79.38	5.50	4.70	
Proportion with large OOP expenses.	0.43	4.82	3.75	
Medicaid recipients				
Visits:				
	0.00	6.63	E ao	
Hospital	0.33	6.63	5.20	
Facility	1.36	7.70	6.27	
Office	5.21	5.59	4.63	
Charges:				
Hospital	691.56	13.56	10.55	
Facility	78.09	7.45	5.80	
Office	139.60	7.27	6.04	
Expenses:				
Hospital, OOP	36.18	29.97	23.98	
Facility, OOP		29.97	16.19	
Office, OOP	7.39			
	23.10	9.57	7.44	
Proportion with large OOP expenses.	0.11	22.79	18.32	

Table 3.	Estimated means and relative standard errors for the unlinked National Medical Expenditure Survey (NMES) design with 6,000 and
10,000 re	espondent originating base reporting units (OBRU's)—Con.

		Relative st.	andard errors
Domain and outcome measure	Y <sub>k</sub> ( <i>NMES</i> )	6,000-respondent OBRU's	10,000-respondent OBRU's
College head of household population			
/isits:			
Hospital	0.14	7.17	5.72
Facility	0.75	9.91	8.20
Office	4.80	4.33	3.37
Charges:			
Hospital	287.87	19,18	15.06
Facility	45.17	8.66	7.22
Office	141.41	4.84	3.76
xpenses:			
Hospital, OOP	40.34	42.30	32.84
Facility, OOP	8.85	11.22	8.73
Office, OOP	75.15	5.71	4.43
Proportion with large OOP expenses,	0.30	14.48	11.45

Table 4. Su	ummary of estimated co	sts of project tasks for the 6,00	0-respondent originating bas	e reporting unit unlinked design
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		Project task <sup>1</sup>					
·	Cost category	Total	1	2	3	4	5
01	Total	\$4,963,013	\$56,781	\$185,500	\$59,566	\$515,553	\$1,618,746
	Direct technical labor						
02 03	On-site Off-site	1,242,967 292,583	52,732 -	20,078 -	11,943 23,669	30,144 50,291	228,215 197,350
	Other direct cost						
04	Total	3,427,463	4,049	165,422	23,954	435,118	1,193,181
05 06 07	Materials and supplies Services Shipping and communications Travel:	58,565 183,871 162,204	147 311 305	1,034 1,087 646	839 2,163 6,161	473 24,449 11,289	24,222 20,779 74,541
08 09 10	On-site Off-site Consultants	52,365 219,079 27.825	842 - -	1,869 139	1,015 7,695 -	10,479 148,441	8,872 48,649 -
11 12 13	Computer services Reports and reproductions Interviewer services	681,529 166,369 684.042	-	- 159,748 -	- 38 1.604	- - 139.607	۔ 1,406 502,782
14 15	Interviewer expenses	439,080 124,805	-	-	3,718	95,608	331,716 124,805
16 17 18	Clerical labor Clerical labor surcharge	379,626 197,264	458 2	561 249	28 10	1,893 1,088	27,984 15,827
18	Miscellaneous	14,441 36,398	- 1,984	60 29	85 598	152 1,639	1,367 10,231

<sup>1</sup>Legend for project tasks: 1 = Survey sampling. 2 = Instrument and materials development.

3 = Field preparations. 4 = Survey training.

5 = Data collection.

6 = Control system development and production. 7 = Data receipt, editing, and document control. 8 = Data coding operations.

9 = Data entry operations.
10 = Control card development, maintenance, and production.
11 = Summary development, maintenance, and productron.

12 = Other data processing operations.

13 = Database construction.

14 = Counting and listing (costs not incurred on the National Medical Care Utilization and Expenditure Survey).
 15 = Project administration.

Table 4.	Summary of estimated costs of	project tasks for the 6,000-	respondent originating be	ase reporting unit unlinked design—Con.
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				Project ta	sk1—Con.					
6	7	8	9	10	11	12	13	14	15	
\$217,061	\$280,462	\$72,417	\$388,378	\$104,271	\$167,509	\$613,673	\$401,874	\$57,842	\$223,380	01
81,099	42,188	12,925	89 -	40,089 -	60,378 -	302,596 -	163,058 -	7.559 21,273	189,873 -	02 03
135.962	238,274	59,492	388,289	64,182	107,131	311,077	238,816	29,009	33,507	04
952	4,663	1,494	255	5,174	6,854	10,244	65	473	1,675	05
16	6,467	65	126,173	-	23	2,135	-	203	-	06
3	56,677	99	18	225	126	2,885	845	1,683	6,701	07
863	1,490	293	2	839	1,098	4,449	150	-	20,104	08
-	-	-	-	-	-	-	-	14,155	-	09
1,419	-	-		3,971	819	21,380	236	-	-	10
119,932	•	-	-	49,067	88,360	186,789	237,381	-	-	11
-	-	-	-	-	-	-	-	151	5,026	12
-	-	-	-	-	-	33,436	-	6,613	-	13
-	-	-	-	•	2,307	-	-	5,731	-	14
•	•	•	-	-	-		-	-	-	15
38	127,612	44,693	140,645	13	72	35,574	55	-	-	16
•	40,072	12,179	115,406	3	35	12,364	29	-	-	17
12.329	108	8	169	-	149	14	-	-	-	18 19
410	1,185	661	5,621	4,890	7,288	1,807	55	-	-	12

#### Table 5. Summary of estimated costs of project tasks for the 10,000-respondent originating base reporting unlinked design

		Project task <sup>1</sup>							
	Cost category	Total	1	2	3	4	5		
01	Total	\$7,209,409	\$75,137	\$293,470	\$59,566	\$715,400	\$2,396,714		
	Direct technical labor								
02 03	On-site Off-site	1,591,977 370,071	69,974 -	21,294 -	11,943 23,669	36,747 61,415	288,800 249,532		
	Other direct cost								
04	Total	5,247,360	5,163	272,176	23,954	617,238	1,858,382		
05 06 07	Materials and supplies	78,334 285,138 223,165	166 383 421	1,097 1,153 685	839 2,163 6,161	579 30,013 14,319	30,806 28,417 94,459		
08 09 10	On-site Off-site Consultants	57,606 254,534 27,919	1,177	1,983 147	1,015 7,695	14,253 161,462	8,872 61,639		
11 12 13	Computer services, Reports and reproductions, Interviewer services	1,004,795 273,604 1,113,807	-	266,158	- 38 1.604	- - 231.701	۔ 1,891 818.041		
13 14 15	Interviewer services Interviewer expenses Respondent incentives	708,756 207,938	-	-	3,718	159,134	532,686 207,938		
16 17 18	Clerical labor Clerical labor surcharge	616,264 321,134	517 2	595 264	28 10 85	2,273 1,324	37,662 21,305		
18 19	Miscellaneous Overtime expenses	22,038 52,328	- 2,497	64 30	598	182 1,998	1,733 12,933		

<sup>1</sup>Legend for project tasks: 1 = Survey sampling. 2 = Instrument and materials development.

3 = Field preparations.

4 = Survey training.

5 = Data collection.

6 = Control system development and production.

7 = Data receipt, editing, and document control.
 8 = Data coding operations.

9 = Data entry operations.
10 = Control card development, maintenance, and production.
11 = Summary development, maintenance, and production.
12 = Other data processing operations.
13 = Database construction.
14 = Counting and listing (costs not incurred on the National Medical Care Utilization and Expenditure Survey).
15 = Project administration.

Table 5.	Summary of estimated costs o	project tasks for the 1	,000-respondent originating base reporting unit unlinked design—Cor
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					1- <i>Con.</i>	Project task				
	15	14	13	12	11	10	9	8	7	6
8 (	\$234,018	\$96,403	\$562,075	\$861,801	\$251,632	\$146,186	\$646,941	\$115,142	\$438,640	\$316,284
5 (	198,915 -	12,599 35,455	228,041 -	393,203 -	<sup>.</sup> 87,929 -	56,189 -	147 -	16,785 -	52,116 -	117,295 -
3 (	35,103	48,349	334,034	468,598	163,703	89,997	646,794	98,357	386,524	198,989
	1,755	789	90	15,790	9,802	7,252	417	1,940	5,809	1,203
- (	•	339	-	3,336	36	-	210,189	91	8,992	26
1 (	7,021	2,804	1,181	2,885	126	225	30	164	92,681	3
2 (	21,062	-	209	4,449	1,098	839	3	293	1,490	863
- (	-	23,591	-	-	-	-	-	-	•	-
- 1	-	-	330	21,380	819	3,971	-	-	-	1,419
- :	-	-	332,033	290,764	136,238	70,169	-	-	-	175,591
	5,265	252	-	-	-	-	-	•	•	-
-	-	11,021	-	51,440	-	-	-	-	-	-
- :	-	9,552	-	-	3,666	-	-	-	-	-
- 1	-	-	-	-	-	-	-	-	-	-
- :	-	-	75	56,219	90	<u>,</u> 17	234,351	74,462	209,926	49
- '	-	-	40	19,552	46	4	192,302	20,293	65,992	-
- '	-	-		21	208	-	279	13	134	19,319
- '	-	-	76	2,762	11,574	7,520	9,223	1,101	1,500	516

Table 6. Overview of Research Triangle Institute (RTI) and National Opinion Research Center (NORC) actual National Medical Care Utilization and Expenditure Survey (NMCUES) Household Survey direct cost experience compared with a 1980 NMCUES RTI-only design

Task description	NORC direct cost	RTI direct cost <sup>1</sup>	Consolidation of RTI and NORC direct cost experience	Estimated costs to conduct 1980 RTI- only design	Difference between RTI and NORC actual versus RTI- only design	Percent difference
Total	\$3,194,209	\$3,184,396	\$6,378,605	\$5,220,947	-\$1,157,658	-19
Instrument development	116,106	24,644	140,750	25,857	-114,893	-82
GPO printing (NORC only)	133,565	-	133,565	166,883	33.318	-25
Sampling	61,822	44,291	106,113	58,147	-47.966	-45
HHS data collection	1,163,065	1,006,363	2,169,428	1,814,496	-354,932	-16
HHS training	444,678	433,958	878,636	602,922	-275,714	31
Receipt and editing	156,057	160,852	316,909	290,655	-26.254	8
Coding	64,368	42,066	106,434	75,280	-31,154	-29
Data entry	198,105	204,731	402,836	405,725	2,889	1
Control system production	135,147	146,581	281,728	223,711	58,017	-21
Control card production	96,507	74,502	171,009	107,080	-63,929	37
Summary production	80,797	107,851	188,648	173,175	-15,473	-8
Other data processing	408,940	437,529	846,469	630,352	-216,117	-26
Database construction (RTI only)	-	288,285	288,285	412.646	124,361	43
Project management	135,052	212,743	347,795	234,018	-113,777	-33

<sup>1</sup>RTI task cost experience already ratio adjusted for the National Household Survey.

# Chapter 3 The linked dwelling unit design

The first National Health Interview Survey (NHIS) linked design investigated was a linked dwelling unit design. Using this design, the National Medical Expenditure Survey (NMES) is selected from a frame of NHIS sample listings. For comparison, four design options were developed based on two primary sampling unit (PSU) counts and two sample sizes. The variances achieved by the linked dwelling unit design were modeled and compared with those achieved by the unlinked designs. The sample sizes for the 100 and 200 PSU designs were set so that the resulting samples have the same precision as that of the unlinked originating base reporting unit (OBRU) design with 6,000 and 10,000 responding. Costs were developed for these four design options.

#### Definition

Linkage of NMES to NHIS makes available a list frame of names and addresses for NMES sample selection. The sample units in this design are the addresses included in NHIS rather than NHIS sample persons living at the addresses. After selecting a sample of addresses from the NHIS frame, NMES interviews the occupants of the sample dwelling units in round 1. NHIS sample members who move before round 1 of NMES are not followed; instead, any new occupants of the dwelling are included in NMES. Except for the selection process of the round 1 sample, the linked dwelling unit design follows the same procedures and definitions as those of the unlinked NMES. That is, the first, second, and fifth data collection rounds are conducted by personal interview; the third and fourth rounds, by telephone. Family members who are college students living away from home are interviewed at their temporary addresses. The round 1 sample individuals have data collected for them for the remaining four rounds of the survey whether or not they continue living in the same dwelling.

Using the NHIS listings for NMES sample selection, it was considered whether units that were nonresidential or nonresponding should be excluded before selection of the NMES sample. Units used for nonresidential purposes only would likely be nonresidential at the time of NMES. However, during the time between NHIS and NMES, the use of a nonresidential structure could change or residential spaces could be added. Also, the NHIS interviewer might fail to note a residential apartment attached to a nonresidential unit. These examples suggest that undercoverage in the NMES sample is likely if NHIS-identified nonresidential structures are omitted.

The second consideration for NMES sample selection is

whether to exclude residential listings for which NHIS could not obtain a response. Although the NHIS refusal rate is very low, approximately 2.5 percent, the short data collection period (2 weeks) results in more nonresponse due to absence than in NMES collection (2.5 percent versus 0.6 percent of the 1980 NMCUES). Also, some of these nonresponding households may move before round 1 of NMES and be replaced by more cooperative households. The response rate from new occupants is assumed to be the same as that of the general population. If all nonresponding units were removed from the NMES frame, NMES would start with a 5.0 percent nonresponse rate before data collection and with the associated nonresponse bias.

Because nonresponding and ineligible NHIS listings are likely to yield few responding NMES cases, but excluding them would result in undercoverage of the NMES sample, the best approach is to include them in the frame but sample them at a lower rate. The low cost of identifying a nonresidential unit makes it feasible to include all nonresidential addresses in the frame to avoid undercoverage of the NMES frame. Nonresponding units are also included but the NHIS experience is used to determine the extent of followup for nonresponding units.

Therefore, the frame for NMES should include all of the NHIS sample addresses associated with the NMES sample PSU's and segments. After selection of the round 1 sample addresses, the collection procedures are the same as those of the unlinked design. These include the use of the half-open interval procedure for new construction to be included in NMES.

#### Sample size determination

To compare the linked designs with the unlinked designs, the sample size for the linked designs was set to the size yielding the same precision as the unlinked design. To determine the sample size for the linked dwelling unit design, the variance for the design was modeled.

The redesigned NHIS has the same target population as NMES. To represent this target population, NHIS includes 200 sample PSU's and 8,750 segments from these PSU's. The segments contain an average of 40 addresses, 6 of which are selected for inclusion in NHIS. The sample segments are separated into 52 weekly sets, so that each weekly sample is a valid national sample. A feature of NHIS is that the black population is oversampled at a rate 1.4 times that of all other races. To model the variance of NMES sample estimates, it is assumed that NHIS oversamples black persons by increasing selection of high concentration black segments. To produce a self-weighting NMES, the effect of this oversampling is removed by subsampling these segments. The estimation procedures are similar to those presented for the unlinked design. That is, the sample estimate of mean utilization or expenditure per person-year is estimated by means of equation (1) as

$$\overline{Y}_{k}(\text{NMES}) = \frac{\sum_{i \in S} W(i)\delta_{k}(i)Y(i)}{\sum_{i \in S} W(i)T(i)\delta_{k}(i)}$$
(11)

and the proportion burdened with large out-of-pocket expenditures by means of equation (2) as

$$\overline{Y}_{k}(\text{NMES}) = \frac{\sum_{i \in S} W(i)T(i)\delta_{k}(i)Y(i)}{\sum_{i \in S} W(i)T(i)\delta_{k}(i)}$$
(12)

Using this approach, the variance of  $\overline{Y}_k(\text{NMES})$  can again be expressed as

$$\operatorname{Var}\left[\overline{Y}_{k}(\operatorname{NMES})\right] = \frac{\sigma_{k}^{2}(\operatorname{PSU})}{r} + \frac{\sigma_{k}^{2}(\operatorname{SEG})}{r\overline{s}} + \frac{\sigma_{k}^{2}(\operatorname{OBRU})}{r\overline{s}\overline{t}} \quad (13)$$

- where  $\sigma_k^2(\text{PSU}) =$  between NHIS PSU, within NHIS segment variance component for domain k
  - r = number of NHIS PSU's from which NMES is selected
  - $\sigma_k^2(SEG) =$  between NHIS segment, within NHIS PSU variance component for domain k
    - $\overline{s}$  = average number of NHIS segments selected for NMES per sampled PSU
  - $\sigma_k^2(\text{OBRU}) = \text{between NMES OBRU, within NHIS seg$  $ment variance component for domain k}$ 
    - $\overline{t}$  = average number of NHIS addresses selected for NMES per sampled segment

The specifications for the redesigned NHIS indicate that the NHIS PSU's and segments are similar in definition and size to those of the 1980 NMCUES. For this reason, the 1980 NMCUES variance component estimates described earlier were used to model the NHIS variance components.

The parameters remaining to be specified are r,  $\bar{s}$ , and  $\bar{t}$ . Depending on the design being modeled, the number of PSU's or r is 100 or 200. The NHIS samples 6 addresses out of 40 in a segment. NMCUES data were used to determine the number of responding OBRU's that could be derived from these six addresses. On the average, NMCUES obtained 1.045 responding OBRU's per address. With the same response and attrition rates for the linked design, six responding OBRU's is the maximum that could be obtained per sample segment. Because this is smaller than the optimal number of OBRU's to select per segment, it is assumed that all NHIS sample addresses within NMES-subsampled segments are included in NMES so that  $\bar{t} = 6$ .

The total sample size is  $r\bar{s}\bar{t}$ ; the term remaining to be specified is  $\bar{s}$ . This process is illustrated for the 100-PSU design set to achieve the same precision as that of the unlinked OBRU design with 6,000 responding. The variance for the unlinked NMES OBRU design with 6,000 responding is modeled as

$$\frac{\sigma_k^2(\text{PSU})}{102} + \frac{\sigma_k^2(\text{SEG})}{750} + \frac{\sigma_k^2(\text{OBRU})}{6,000}$$

and the variance for the 100-PSU 6,000-OBRU-equivalent linked design is modeled as

$$\frac{\sigma_k^2(\text{PSU})}{100} + \frac{\sigma_k^2(\text{SEG})}{100\overline{s}} + \frac{\sigma_k^2(\text{OBRU})}{600\overline{s}}$$

These two expressions can be set equal for a specific domain k and a specific statistic, and the value of  $\overline{s}$  derived will result in the linked design achieving the same precision as that of the unlinked design. The required number of segments vary depending on the domain and the outcome measure. Therefore, an average over the 50 statistics formed by the 5 domains and 10 outcome measures was used to determine the number of segments to be costed. Table 7 presents the number of segments required to obtain the precision of the unlinked design for each of the 5 domains and 10 outcome measures.

### Cost modeling

The difference between the linked dwelling unit design and the unlinked design is the selection procedure for sample dwelling units which may affect the response rates for the survey. For example, interviewing the occupants of the sample dwelling units (except for new occupants), who have already been interviewed once, might have a negative effect on response. However, lead letters can be sent before the NMES interview. Because the use of lead letters tends to improve response, the linked dwelling unit design should be able to achieve the same response rates as the unlinked design.

Costs were developed for four linked dwelling unit designs based on the two PSU size options and the two sample size options. These four designs are as follows:

- Design A. 100 PSU's and a sample size sufficient to yield estimates of the same precision as the unlinked design with 6,000 responding OBRU's.
- Design B. 200 PSU's and a sample size sufficient to yield estimates of the same precision as the unlinked design with 6,000 responding OBRU's.
- Design C. 100 PSU's and a sample size sufficient to yield estimates of the same precision as the unlinked design with 10,000 responding OBRU's.
- Design D. 200 PSU's and a sample size sufficient to yield estimates of the same precision as the unlinked design with 10,000 responding OBRU's.

Based on the procedures discussed in the previous section, the sample sizes for the four designs were determined. Design A has 100 PSU's, 976 segments, and 5,856 responding OBRU's; design B has 200 PSU's, 921 segments, and 5,526 responding OBRU's; design C has 100 PSU's, 1,629 segments, and 9,774 responding OBRU's; and design D has 200 PSU's, 1,489 segments, and 8,934 responding OBRU's.

For each of these designs, all sample addresses are visited regardless of their classification by NHIS. Because the response rates are assumed to be the same as those of the 1980 NMCUES, the unit costs for the linked dwelling unit design are similar to those of the unlinked design. Costs for lead letters were added to the model, and the costs for counting and listing were deleted from the model.

Using these unit costs, the direct costs were estimated for the four designs. These costs are summarized in tables 8-11. The total costs for all tasks and all data collection rounds were 4,871,106 for design A and 4,947,848 for design B. For the equivalent 6,000-OBRU unlinked design, the total cost was 4,963,013. The costs for design A are less due to not having counting and listing costs and sampling 100 instead of the 102 PSU's in the unlinked design. Design B is more costly because it samples 200 PSU's. The direct cost estimate for designs C and D are 7,147,752 and 6,930,673, respectively, compared with 7,209,409 for the equivalent OBRU unlinked design with 10,000 responding. Both designs have costs lower than those of the unlinked design, and the 200-PSU design has the lowest total cost. This suggests that increased precision constraints make it cost effective to increase the number of PSU's in the design to 200. For reasons described in chapter 5, these results, instead, appear to be an indication of instability in the variance component estimates.

#### Other design considerations

The linked dwelling unit design, as described in this chapter, makes little use of the information collected for NHIS respondents. An alternative approach is to stratify NHIS dwelling units based on the characteristics of the occupants. Strata are also developed for the units that were unoccupied, nonresidential, and nonresponding. This stratification might improve the efficiency of the designs described earlier. Such an approach involves an optimization to determine the appropriate sample sizes. Optimization requires modeling the effect of movement on stratification. Depending on the amount of movement, there may be no advantage in stratifying the NHIS addresses before the NMES sample selection. Because of the complexity of the variance modeling and the assumption that the advantage of stratification is small as a result of movement, the stratification approach was not investigated in this study.

#### Table 7. Required segment size for the linked design to obtain the precision of the unlinked design by domain and type of service

	Precision					
	6,000-respo	ndent OBRU's	10,000-respondent OBRU'			
Domain and outcome measure	100 PSU's	200 PSU's	100 PSU's	200 PSU		
Total						
Visits:						
Hospital	1,006	849	1,684	1,287		
Facility	934	718	1,569	1,041		
Office	971	826	1,626	1,257		
Charges:	005	000	1.050	1 640		
Hospital	995 950	988 831	1,658 1,590	1,640 1,282		
Facility	947	941	1,550	1,202		
Expenses:	347	541	1,570	1,502		
Hospital, out-of-pocket (OOP)	988	982	1,646	1,630		
Facility, OOP	988	868	1,653	1,342		
Office, 00P	913	908	1,521	1,510		
Proportion with large OOP expenses	916	912	1,528	1,515		
	010	012	1,020	.,		
150 percent of poverty population						
Visits:						
Hospital	979	973	1,631	1,615		
Facility	920	916	1,533	1,522		
Office	953	949	1,589	1,576		
Charges:	976	971	1 007	1 010		
Hospital	976	926	1,627 1,552	1,612 1,537		
Office	960	848	1,606	1,316		
Expenses:	300	040	1,000	1,310		
Hospital, OOP.	1,000	994	1,667	1.650		
Facility, OOP	979	973	1,632	1,614		
Office, OOP.	954	949	1,590	1,576		
Proportion with large OOP expenses	958	953	1,597	1,583		
Medicare recipients						
Visits:						
Hospital	1,004	898	1,679	1,402		
Facility	1,000	992	1,667	1,644		
Office	1,013	750	1,704	1,071		
Charges:						
Hospital	1,003	907	1,677	1,424		
Facility	1,000	992	1,667	1,644		
Office	1,003	807	1,682	1,196		
Expenses:	1 000	002	1 667	1 649		
	1,000	993	1,667	1,648		
Facility, OOP	1,000	976 777	1,668 1,635	1,603 1,146		
Proportion with large OOP expenses	975 975	969	1,626	1,609		
Medicaid recipients						
Visits:						
Hospital	987	968	1,646	1,594		
Facility	946	861	1,581	1,356		
Office	995	870	1,664	1,341		
Charges:						
Hospital	985	978	1,641	1,624		
Facility	973	966	1,621	1,603		
Office	1,005	873	1,682	1,343		
Expenses:						
Hospital, OOP	1,002	947	1,672	1,526		
Facility, OOP	1,000	992	1,667	1,644		
••						
Office, OOP Proportion with large OOP Expenses	996 966	991 906	1,661 1,614	1,645 1,451		

See note at end of table.

Table 7.	Required segment size for the linked dea	ion to obtain the precision of the unlinked	design by domain and type of service—Con.

	Precision					
	6,000-respo	ndent OBRU's	10,000-respondent OBRU's			
Domain and outcome measure	100 PSU's	200 PSU's	100 PSU's	200 PSU's		
College head of household population						
Visits:						
Hospital	998	950	1,666	1,535		
Facility	951	836	1,590	1,293		
Office	972	967	1,621	1,606		
Charges:						
Hospital	987	966	1,645	1,588		
Facility	1,006	866	1,683	1,324		
Office	969	963	1,615	1,599		
Expenses:						
Hospital, OOP	978	975	1,630	1,621		
Facility, OOP	955	949	1,593	1,576		
Office, OOP	947	942	1,578	1,566		
Proportion with large OOP expenses	974	943	1,625	1,539		
Average						
All outcome measures	976	921	1,629	1,489		

NOTE: OBRU's = originating base reporting units; PSU's = primary sampling units.

#### Table 8. Summary of estimated costs of project tasks for linked dwelling unit design A

		Project task <sup>1</sup>						
	Cost category	Total	1	2	3	4	5	
01	Total	\$4,871,106	\$55,462	\$182,510	\$52,723	\$560,811	\$1,587,592	
	Direct technical labor							
02 03	On-site Off-site	1,230,433 274,965	49,575 -	20,076	13,164 21,278	30,411 50,741	232,589 202,946	
	Other direct cost							
04	Total	3,365,708	5,887	162,434	18,281	479,659	1,152,057	
05 06	Materials and supplies	58,681 180,106	161 327	1,051 1,076	658 1,582	475 24,673	25,175 20,858	
07	Shipping and communications Travel:	160,572	308	645	4,971	10,245	77,561	
08 09	On-site	52,738 242,726	834 -	1,872 139	750 6,586	11,796 186,203	8,201 49,798	
10 11	Consultants	27,821 669,623	-	-	-	-	-	
12	Reports and reproductions	163,302	-	156,734	12	-	- 1,530	
13	Interviewer services	653,241	-	-	608	143,303	476,550	
14	Interviewer expenses	415,764	•	-	2,514	98,151	312,841	
15	Respondent incentives	121,771	-	•	-	-	121,771	
16	Clerical labor	374,985	1,931	581	20	1,908	29,442	
17	Clerical labor surcharge	194,182	361	248	7	1,098	16,544	
18		14,166	-	60	67	153	1,374	
19	Overtime expenses	36,029	1,965	28	506	1,654	10,412	

<sup>1</sup>Legend for project tasks:

1 = Survey sampling.

2 = Instrument and materials development.

3 = Field preparations.

4 = Survey training.

5 = Data collection.

6 = Control system development and production.

7 = Data receipt, editing, and document control.

8 = Data coding operations.
9 = Data entry operations.
10 = Control card development, maintenance, and production.

11 = Summary development, maintenance, and production.
12 = Other data processing operations.
13 = Database construction.

14 = Project administration.

Table 8. Summary of estimated costs of project tasks for linked dwelling unit design A-Con.

				ın.	Project task <sup>1</sup> —Co	F			
	14	13	12	11	10	9	8	7	6
0.	\$223,380	\$395,963	\$604,526	\$164,423	\$102,712	\$378,794	\$70,829	\$278,007	\$213,374
02 03	189,873 -	160,658 -	299,240 -	59,363 -	39,491 -	87 -	12,782 -	42,383 -	79,741 0
04	33,507	235,305	305,286	105,060	63,221	378,707	58,047	234,624	133,633
05 06 07	1,675 - 6,701	63 - 832	10,043 2,089 2,885	6,745 22 126	5,092 - 225	247 123,031 18	1,476 63 96	4,876 6,369 55,956	944 16 3
08	20,104	147	4,449	1,098	839	2	293	1,490	863
09 10 11	-	- 232 233,895	- 21,380 182,966	- 819 86,609	- 3,971 48,282	-	-	-	- 1,419 117,871
12 13	5,026	-	32,780			-	-	-	-
14 1!	-	-		2,258	-	•	-	-	-
1	-	53 29	34,811 12,099	70 35	13 4	137,183 112,571	43,588 11,879	125,347 39,307	38
18 19	-	- 54	13 1,771	146 7,132	- 4,795	165 5,490	7 645	107 1,172	12,074 405

#### Table 9. Summary of estimated costs of project tasks for linked dwelling unit design B

		Project task <sup>1</sup>						
<u></u>	Cost category	Total	1	2	3	4	5	
01	Total	\$4,947,848	\$54,010	\$173,550	\$93,672	\$535,022	\$1,742,799	
	Direct technical labor							
02 03	On-site	1,268,244 334,853	48,311 -	19,974 -	22,397 37,922	29,796 49,707	281,442 247,224	
	Other direct cost							
04	Total	3,344,751	5,699	153,576	33,353	455,519	1,214,133	
05	Materials and supplies	62,026	154	1,045	1,152	465	28,895	
06	Services	176,378	316	1,071	3,163	24,160	23,146	
07	Shipping and communications Travel:	177,169	297	642	8,742	9,857	92,367	
08	On-site	57,969	816	1,862	1.302	11.558	13,151	
09	Off-site	248,748	-	138	11,659	177,414	59,537	
10	Consultants	27,814	-	-	-		-	
11	Computer services	642,969	-	-	-	-	-	
12	Reports and reproductions	155,062	-	147,906	24	•	2,106	
13	Interviewer services	644,907	-	-	1,217	134,931	477,463	
14	Interviewer expenses	423,965	-	-	5,027	92,416	324,376	
15	Respondent incentives	114,913	-	-	-	-	114,913	
16	Clerical labor	369,841	1,848	577	40	1,872	40,883	
17	Clerical labor surcharge	191,442	345	247	14	1,076	22,944	
18	Miscellaneous	13,892	-	60	118	150	1,640	
19	Overtime expenses	37,655	1,923	28	895	1,620	12,712	

<sup>1</sup>Legend for project tasks:

1 = Survey sampling.

2 = Instrument and materials development. 3 = Field preparations.

4 = Survey training.

5 = Data collection.

6 = Control system development and production.

7 = Data receipt, editing, and document control.

8 = Data coding operations.

9 = Data entry operations.

10 = Control card development, maintenance, and production.

11 = Summary development, maintenance, and production.

12 = Other data processing operations. 13 = Database construction.

14 = Project administration.

Table 9. Summary of estimated c	osts of project tasks for linked dwelling unit design B—Con.
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		Project task <sup>1</sup> —Con.									
	14	13	12	11	10	9	8	7	6		
01	\$223,380	\$382,754	\$584,065	\$157,488	\$99,255	\$357,472	\$67,306	\$271,883	\$205,192		
02 03	189,873 -	155,299 -	291,767 -	57.092 -	38,163 -	83	12,464 -	44,828 -	76,755 -		
04	33,507	227,455	292,298	100,396	61,092	357,389	54,842	227,055	128,437		
05 06 07	1,675 - 6,701	61 - 804	9,586 1,990 2,885	6,502 21 126	4,921 - 225	233 116,102 17	1,439 61 91	4,975 6,332 54,412	923 16 3		
08	20,104	142	4,449	1,098	839	2	293	1,490	863		
09 10 11	-	- 225 226,091	- 21,380 174,394	- 819 82,662	- 3,971 46,541	-	-	-	- 1,419 113,281		
12	5,026	-	31,296	-		-	-	-	-		
14 15	-	-		2,146	-	-	-	-	•		
16 17	-	52 28	33,108 11,506	69 34	13 4	129,456 106,230	41,133 11,210	120,753 37,804	37		
18 19	-	- 52	12 1,692	141 6,778	- 4,578	156 5,193	7 608	110 1,179	11,498 397		

#### Table 10. Summary of estimated costs of project tasks for linked dwelling unit design C

		Project task <sup>1</sup>						
	Cost category	Total	1	2	3	4	5	
01	Total	\$7,147,752	\$72,697	\$288,961	\$52,723	\$856,020	\$2,359,208	
	Direct technical labor							
02 03	On-site	1,574,037 344,486	64,592 -	21,296 -	13,164 21,278	37,460 62,579	298,543 260,529	
	Other direct cost							
04	Totai	5,229,229	8,105	267,665	18,281	755,981	1,800,036	
05	Materials and supplies	79,277	241	1,127	658	587	32,799	
06	Services	280,203	455	1,134	1,582	30,541	28,798	
07	Shipping and communications Travel:	225,072	440	684	4,971	14,679	100,307	
08	On-site	56,813	1,043	1,986	750	14,530	8,201	
09	Off-site	357,644	-	147	6,586	286,792	64,119	
10	Consultants	27,914	-	-	-	-	-	
11	Computer services	986,252	-	-	-	-	-	
12	Reports and reproductions	269,012	-	261,601	12	-	2,134	
13	Interviewer services	1,067,987	-	-	608	239,148	777,819	
14	Interviewer expenses	673,587	-	-	2,514	163,797	503,687	
15	Respondent incentives	203,245	-	-	-	-	203,245	
16	Clerical labor	611,365	2,914	629	20	2,326	40,843	
17	Clerical labor surcharge	317,242	544	263	7	1,356	22,969	
18	Miscellaneous	21,642	-	64	67	187	1,762	
19	Overtime expenses	51,974	2,468	30	506	2,038	13,353	

<sup>1</sup>Legend for project tasks:

1 = Survey sampling.

2 = Instrument and materials development.

3 = Field preparations. 4 = Survey training.

5 = Data collection.

6 = Control system development and production. 7 = Data receipt, editing, and document control.

8 = Data coding operations.

9 = Data entry operations.
 10 = Control card development, maintenance, and production.

11 = Summary development, maintenance, and production.

12 = Other data processing operations.

13 = Database construction. 14 = Project administration.

Table 10.	Summary of estimated costs of project tasks for linked dwelling unit design C-Con.
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				n.	roject task <sup>1</sup> —Co	Pi			
	14	13	, 12	11	10	9	8	7	6
01	\$234,018	\$552,884	\$847,568	\$246,811	\$143,776	\$632,080	\$112,686	\$437,744	\$310,576
02 03	\$198,915 -	224,311 -	387,998 -	86,349 -	55,263 -	143 -	16,564 -	54,234 -	115,205 -
04	35,103	328,573	459,570	160,462	88,513	631,937	96,122	383,510	195,371
05 06 07	1,755 - 7,021	88 - 1,162	15,473 3,267 2,885	9,631 35 126	7,131 - 225	407 205,347 30	1,913 89 161	6,277 8,929 92,378	1,190 26 3
08 09	21,062	206	4,449	1,098	839	3	293	1,490	863
10 11	-	325 326,602	21,380 284,804	819 113,497	3,971 68,955	-	-	-	1,419 172,394
12 13	5,265 -	-	- 50,412	-	-	-	-	-	-
14 15 16	-	- - 75	- - 55.034	3,589 - 88	- - 17	- - 228,970	- - 72,751	- - 207,650	- - 48
17 18	-	40	19,139 20	46 204	5	187,891 272	19,827 12	65,155 135	- 18,919
19	•	75	2,707	11,329	7,370	9,017	1,076	1,496	509

Table 11. Summary of estimated costs of project tasks for linked dwelling uni-
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			Project task <sup>1</sup>						
	Cost category	Total	1	2	3	4	5		
01	Total	\$6,930,673	\$69,003	\$266,142	\$93,672	\$792,966	\$2,413,978		
	Direct technical labor								
02 03	On-site	1,568,321 395,372	61,373 -	21,034 -	22,397 37,922	35,954 60,050	338,811 297,400		
	Other direct cost								
04	Total	4,966,980	7,630	245,108	33,353	696,962	1,777,767		
05	Materials and supplies	79,954	224	1,111	1,152	563	35,527		
06	Services	263,471	428	1,122	3,163	29,288	30,054		
07	Shipping and communications Travel:	233,331	412	676	8,742	13,732	112,152		
08	On-site	61,649	998	1,962	1,382	13,946	13,151		
09	Off-site	349,105	-	145	11,659	265,307	71,994		
10	Consultants	27,894	-	-	-	-	-		
11	Computer services	918,373	-	-	-	-	-		
12	Reports and reproductions	247,043	-	239,122	24	-	2,632		
13	Interviewer services	1,006,041	-	-	1,217	218,676	739,516		
14	Interviewer expenses ,,	648,487	-		5,027	149,776	490,381		
15	Respondent incentives	185,781	-	-	-	-	185,781		
16	Clerical labor	575,451	2,703	618	40	2,237	50,800		
17	Clerical labor surcharge	298,481	505	260	14	1,301	28,532		
18	Miscellaneous	20,394	-	63	118	180	1,977		
19	Overtime expenses	51,525	2,360	29	895	1,956	15,270		

<sup>1</sup>Legend for project tasks:
1 = Survey sampling.
2 = Instrument and materials development.
3 = Field preparations.
4 = Survey training.
5 = Data collection.
6 = Control system development and production.
7 = Data receipt editing and document control 7 = Data receipt, editing, and document control.

8 = Data coding operations.
9 = Data entry operations.
10 = Control card development, maintenance, and production.
11 = Summary development, maintenance, and production.
12 = Other data processing operations.
13 = Database construction.
14 = Project administration.

Table 11. Summary of estimated costs of project tasks for linked dwelling unit design D-Con.

	Project task <sup>1</sup> Con.								
	14	13	12	11	10	9	8	7	6
01	\$234,018	\$519,241	\$795,467	\$229,146	\$134,972	\$577,786	\$103,712	\$410,830	\$289,740
02 03	198,915 -	210,057 -	368,971 -	80,563 -	51,881 -	131	15,753 -	54,267 -	107 <i>.</i> 604 -
04	35,103	308,574	426,496	148,583	83,091	577,655	87,959	356,563	182,136
05 06 07	1,755 - 7,021	83 - 1,091	14,309 3,014 2,885	9,013 32 126	6,694 - 225	373 187,704 27	1,819 83 147	6,194 8,559 86,092	1,137 24 3
08 09 10	21,062 - -	193 - 305	4,449 - 21,380	1,098 - 819	839 - 3,971	3 - -	293 - -	1,490 - -	863 - 1,419
11 12 13	- 5,265 -	306,724 - -	262,975 - 46,632	123,444	64,524 - -	- - -	- - -	• - -	160,706 - -
14 15 16 17	-	- 70 38	- - 50,699 17,629	3,303 - 84 43	- - 16 4	- - 209,294 171,744	- - 66,500 18,123	- 192,344 60,288	46
18 19	-	70	18 2,506	192 10,429	6,818	249 8,261	11 983	135 1,461	17,451 487

## Chapter 4 The linked household design

Another approach to linking the National Medical Expenditure Survey (NMES) to the National Health Interview Survey (NHIS) is to designate as sampling units the NHIS sample households rather than the sample addresses. This approach facilitates data collection because sample members are known in advance. However, some sample members will move before round 1 and will have to be located. This approach was investigated using the two primary sampling unit (PSU) size options and the two precision constraint sets of the originating base reporting unit (OBRU) unlinked designs with 6,000 and 10,000 responding.

#### Definition

The linked household design selects NHIS households rather than dwelling units. However, the sampling units are the individual members of these subsampled NHIS households. These individuals are key members of the NMES sample. These key individuals are interviewed in round 1 of NMES whether or not they live at the same NHIS address. Thus, tracing and followup of movers is needed in the first round of data collection. Because family-level analyses are conducted in NMES, the members of families formed by the sample individuals need to be interviewed. Most households remain the same in the time period between NHIS and NMES. Because individuals within NHIS households are selected as a group, stable households are entirely composed of NMES key individuals.

Movement into and out of established families is not uncommon, however. The guidelines for handling this movement in round 1 are similar to those used in later rounds of NMES under all design options. That is, individuals who join families formed by key individuals through birth or return from the military, an institution, or overseas residence are included as key individuals in NMES. Other individuals joining the families of key individuals are classified as nonkey. The distinction between key and nonkey sample members is that only key individuals are included in person-level analyses. Data for nonkey persons are only used in developing family-level aggregates. Key individuals are followed through all five rounds of data collection. Nonkey individuals have data collected only for the time period in which they belong to a family containing a key individual.

The frame for the linked household design is a list of NHIS sample households with names, addresses, and information needed for tracing. NHIS not-at-home cases are also included but not NHIS refusals. The frame is stratified based on characteristics related to NHIS oversampling to produce a selfweighting sample.

Because the short NHIS data collection period results in a large percent of nonresponse due to failure to find someone at home, excluding these cases would adversely affect the NMES response rate. Including these addresses presents a problem, however, because residents present at the time of the NHIS interview may move prior to the NMES round 1 interview and be replaced by new tenants. The movement problem can be handled by including special screening procedures for NHIS not-at-home cases. However, the problems associated with movement from NHIS refusals led to their exclusion from the frame for this design.

#### Sample size determination

In a procedure similar to that discussed in the previous chapter, sample sizes were developed for the four designs resulting from the two PSU size options and the two sets of variance constraints. First the design variance was modeled. The intent was not to build an optimal design so only NHIS oversampling was removed and the design was not stratified prior to selection. Therefore, the variance modeling and sample size determination are the same as those described for the linked dwelling unit design. However, converting responding OBRU's into the required number of reporting unit interviews is different from the linked dwelling unit design.

#### **Cost modeling**

The target population for NMES is the civilian noninstitutionalized residents of the United States during the data collection year. Sample individuals are eligible for NMES data collection only during the time they are civilian, noninstitutionalized, and residing in the United States. Determining the costs for NMES required modeling the rate at which NHIS individuals leave the NMES target population through death, institutionalization, or emigration, before the NMES data collection period.

Response and attrition rates differ for the linked household design. Loss occurs due to movement before NMES as well as attrition effects associated with the previous NHIS interview. Tracing is needed in round 1, and more interviews need to be conducted outside the sample clusters, due to the additional movement occurring before round 1.

The first step in the costing process was to model the 1980 NMCUES experience. Movement could only be detected for NMCUES when there was a change of ZIP code. First the ZIP codes associated with the original clustered addresses were determined. In each data collection round, the reporting units (RU's) were classified as to whether the interview occurred within the ZIP-code-defined clustered areas. Additional interviewer travel time and expenses are incurred for interviews outside clustered areas. The only interviews occurring outside the sample clusters in round 1 were for college students living away from home.

When a household moves, there is a one-time only tracing cost to determine the new address. To model this event, a move was defined as when the ZIP code in a round differs from that of the previous round. Both movement outside the clusters and tracing are expected to be greater for the linked household design.

Table 12 presents the results of this modeling of the 1980 NMCUES. Because NMCUES costs occurred to the reporting unit level, these sample sizes are given for RU's. Because the 1980 NMCUES was a clustered area sample of addresses, many of the selections were ineligible units (vacant, nonresident, and so forth), which accounts for the large number of ineligible RU's in round 1. College students living away from home require a separate interview and, thus, are assigned a separate RU number. These college students living away from home account for the 92 RU interviews conducted outside the sample clusters in round 1. By definition, no tracing was needed in round 1. After round 1, there were costs associated with following up sample members who were ineligible or lost to the survey population due to death, institutionalization, entrance into the military, or migration out of the country. There were also costs associated with attempting interviews with nonrespondents. In round 2, for instance, 6,727 RU's were fielded. Of these, 14 were ineligible for the study, 199 failed to respond, and 6,514 completed interviews. Of the 6,514 completing interviews, 395 had moved since round 1, requiring tracing and perhaps a reassignment of the RU to another interviewer. The 6,514 completed interviews had 6,352 conducted within the ZIP code areas associated with the initial sample selections and 162 outside these areas. The 395 RU's requiring tracing may or may not have moved outside the sample clustered ZIP codes. After round 2, these cases did not require additional tracing unless they moved again. However, those of the 395 RU's who moved outside the sample clusters required more interviewer traveltime and expenses to complete their interviews.

The expected sample sizes needed to yield the required number of completed OBRU interviews are given in table 13 for the four linked household designs. Assumptions were made in deriving these sample sizes. First, the required number of responding OBRU's were converted into RU costing units by assuming that the ratio of the number of completed interviews in a round and the number of responding OBRU's would be the same for all designs. With this assumption, the number of completed interviews in each round was estimated as the product of the number of responding OBRU's times each round's ratio of completed RU interviews to responding OBRU's.

Because the linked household design will encounter movement in round 1, the percent of interviews outside the sample clusters should be greater than that in round 1 of the unlinked design. To estimate the extent of the movement, it was assumed that the linked household design encounters similar movement outside the clusters in round 1 to that of the unlinked design in round 2; that round 2 movement outside the clusters is similar to that encountered by the unlinked design in round 3; and so forth. These projected rates were modified to account for less interviewing outside the clusters in round 4 when college students have returned home for the summer. The percent of the completed interviews where tracing is required should be similar in the linked and unlinked designs, except for round 1 of the unlinked design, which does not encounter movement. The round 2 tracing rate for the unlinked design was used to model the round 1 tracing rate for the linked household design.

Modeling the response rate was the next step. The cumulative responses and attrition rates that the 1980 NMCUES encountered were 91.1 percent in round 1; 90.7 percent in round 2; 89.7 percent in round 3; 89.3 percent in round 4; and 89.0 percent in round 5. Excluding the 2.5 percent NHIS refusals from the NMES frame allows the linked household survey better roundwise response rate than that of the unlinked design. The fact that the sample would have been interviewed once already would have a negative effect. Balancing these two factors, the cumulative attrition and response rate expected in the field is 92.5 percent in round 1; 91.5 percent in round 2; 91.1 percent in round 3; 90.8 percent in round 4; and 90.5 percent in round 5. An additional 2.5 percent of the NMES sample would be lost due to NHIS refusal and exclusion from the frame, resulting in effective cumulative response and attrition rates of 90.2, 89.2, 88.8, 88.5, and 88.2 percent in rounds 1 through 5, respectively.

The rate at which sample members become ineligible was modeled in a procedure similar to that of the tracing rate model. That is, it was assumed that in every round after the first the percent ineligible of the total sample fielded is the same for the linked household design as for the 1980 NMCUES. The round 1 ineligible rate for the linked household design was based on the rate in round 2 of the 1980 NMCUES.

Unit costs were developed by round to include identifying ineligible RU's, attempting to interview nonresponding RU's, completing interviews within the sample clusters, completing interviews outside the sample clusters, and tracing movers. These unit costs were used in modeling the costs for the four linked household designs. These costs are presented in tables 14-17. The 6,000-OBRU-equivalent linked household design has direct costs of \$4,891,831 with 100 PSU's and \$4,967,406 with 200 PSU's, compared with \$4,963,013 for the unlinked 6,000-OBRU design. The 10,000-OBRU-equivalent linked household design has direct costs of \$7,182,341 with 100 PSU's and \$6,962,291 with 200 PSU's, compared with \$7,209,409 for the unlinked 10,000-OBRU design. These results suggest that 200 PSU's are more cost efficient for the 10.000-OBRU precision constraints than 100 PSU's, but are more likely a reflection of instability of the variance constraints. (See chapter 5.)

The cost savings associated with linkage are not substantial. Savings for the design, a slightly larger response rate, and no counting and listing costs, are partly offset by added costs associated with tracing movers.

#### Other design considerations

Between the time of the NHIS interview and the beginning of the NMES data collection year, individuals enter the target population through birth or through return from the military, an institution, or overseas. The unlinked household design updates the sample in round 1 using the same procedure as that of all NMES designs. That is, individuals who joined families formed by NMES subsampled individuals enter the survey as key individuals if they were born or returned from an ineligible state after the NHIS interview. This procedure results in undercoverage of the individuals entering the target population who do not join preexisting families. All NMES designs encounter this type of undercoverage in rounds 2-5 of the study, but only the linked household design encounters this in round 1. This undercoverage is not substantial enough to preclude the use of the linked household design, but the dwelling unit design is preferable for optimum population coverage.

By restricting attention to self-weighting designs, thus far, many of the advantages associated with the linkage of NMES to NHIS have been eliminated. The next chapter departs from the self-weighting constraint to investigate optimal versions of the linked household design.

#### Table 12. Sample sizes for the 1980 National Medical Care Utilization and Expenditure Survey (NMCUES) design

			RU's completing interviews			
Round	Reporting units (RU's) ineligible	RU's nonresponding	Tota/	Traced	Inside sample clusters	Outside sample clusters
1	1,115	643	6,601		6,509	92
2	14	199	6,514	395	6,352	162
3	24	94	6,525	248	6,355	170
4	3	72	4,558	114	4,482	76
5	26	57	6,559	183	6,418	141

### Table 13. Sample sizes for the National Medical Expenditure Survey (NMES) linked household design

				RU's compl	eting interview	5
Design type and round	Reporting units (RU's) ineligible	RU's nonresponding	Total	Traced	Inside sample clusters	Outside sample clusters
Design A						
Round 1	14	498	6,165	172	6,012	153
Round 2	23	184	6,084	369	5,925	159
Round 3	23	86	6.094	232	5,935	159
Round 4	3	66	4,257	106	4,186	71
Round 5	24	52	6,126	171	5,981	145
Design B						
Round 1	13	470	5,818	162	5,674	144
Round 2	22	174	5,741	348	5,591	150
Round 3	22	81	5.751	219	5,601	150
Round 4	3	62	4.017	100	3,950	67
Round 5	23	49	5,781	161	5,644	137
Design C						
Round 1	23	831	10,290	287	10,035	255
Round 2	38	307	10,154	616	9,889	265
Round 3	38	144	10,172	387	9,907	265
Round 4	5	110	7,105	177	6,986	119
Round 5	40	87	10,225	285	9,983	242
Design D						
Round 1	21	760	9,406	262	9,173	233
Round 2	35	281	9,282	563	9,039	243
Round 3	35	131	9,297	354	9,054	243
Round 4	5	101	6,495	162	6,387	108
Round 5	37	79	9,346	261	9,125	221

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	Table 14.	Summary of estimated	d costs of project tasks	for linked household design A
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01 02 03 04 05 06 07 08 09 10 11 12 13 14 15		Project task <sup>1</sup>						
	Cost category	Total	1	2	3	4	5	
01	Total	\$4,891,831	\$55,462	\$182,510	\$52,723	\$560,811	\$1,608,317	
	Direct technical labor							
-	On-site Off-site	1,230,862 275,026	49,575 -	20,076 -	13,164 21,278	30,411 50,741	233,018 203,007	
	Other direct cost							
04	Total	3,385,943	5,887	162,434	18,281	479,659	1,172,292	
05	Materials and supplies	58,620	161	1,051	658	475	25,114	
06	Services,	180,229	327	1,076	1,582	24,673	20,981	
07	Shipping and communications	161,436	308	645	4,971	10,245	78,425	
08	On-site	52,738	834	1,872	750	11,796	8,201	
09	Off-site	243,404	-	139	6,586	186,203	50,476	
10	Consultants	27,821	-	-	-	-	-	
11	Computer services	669,623	-	-	-	-	-	
	Reports and reproductions	163,302	-	156,734	12	-	1,530	
	Interviewer services	662,362	-	-	608	143,303	485,671	
	Interviewer expenses	424,103	-	-	2,514	98,151	321,180	
15	Respondent incentives	121,771	-	•	-	-	121,771	
16	Clerical labor	375,848	1,931	581	20	1,908	30,305	
17	Clerical labor surcharge	194,613	361	248	7	1,098	16,975	
18	Miscellaneous	14,166	-	60	67	153	1,374	
19	Overtime expenses	35,906	1,965	28	506	1,654	10,289	

<sup>1</sup>Legend for project tasks: 1 = Survey sampling. 2 = Instrument and materials development.

3 = Field preparations.

4 = Survey training. 5 = Data collection.

6 = Control system development and production.

7 = Data receipt, editing, and document control.

8 = Data coding operations.

 9 = Data entry operations.
 10 = Control card development, maintenance, and production. 11 = Summary development, maintenance, and production.

12 = Other data processing operations.

13 = Database construction.

14 = Project administration.

 Table 14.
 Summary of estimated costs of project tasks for linked household design A—Con.

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	Project task <sup>1</sup> —Con.								
	14	13	12	11	10	9	8	7	6
01	\$223,380	\$395,963	\$604,526	\$164,423	\$102,712	\$378,794	\$70,829	\$278,007	\$213,374
02 03	189,873 -	160,658 -	299,240 -	59,363 -	39,491 -	87 -	12,782	42,383 -	79,741 0
04	33,507	235,305	305,286	105,060	63,221	378,707	58,047	234,624	133,633
05 06 07	1,675 - 6,701	63 - 832	10,043 2,089 2,885	6,745 22 126	5,092 - 225	247 123,031 18	1,476 63 96	4,876 6,369 55,956	944 16 3
08	20,104	147	4,449	1,098	839	2	293	1,490	863
09 10 11	-	- 232 233,895	- 21,380 182,966	- 819 86,609	- 3,971 48,282	-	-	-	- 1,419 117,871
12 13	5,026	-	32,780	-	-	-	-	-	-
14 15	-	-	-	2,258	•	-	-	-	-
16 17	-	53 29	34,811 12,099	70 35	13 4	137,183 112,571	43,588 11,879	125,347 39,307	38
18 19	-	- 54	13 1,771	146 7,132	4,795	165 5,490	7 645	107 1,172	12,074 405

Table 15. Summary of estimated costs of project tasks for linked household des	sign B
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				Proje	ct task <sup>1</sup>	_	
	Cost category	Total	1	2	3	4	5
01	Total	\$4,967,406	\$54,010	\$173,550	\$93,672	\$535,022	\$1,762,357
	Direct technical labor						
02 03	On-site Off-site	1,268,649 334,910	48,311 -	19,974 -	22,397 37,922	29,796 49,707	281,847 247,281
	Other direct cost						
04	Total	3,363,847	5,699	153,576	33,353	455,519	1,233,229
05	Materials and supplies	61,968	154	1,045	1,152	465	28,837
06	Services	176,495	316	1,071	3,163	24,160	23,263
07	Shipping and communications Travel:	177,983	297	642	8,742	9,857	93,181
08	On-site	57,969	816	1,862	1,302	11,558	13,151
09	Off-site	249,388		138	11,659	177,414	60,177
10	Consultants	27,814	-	•	-	-	-
11	Computer services	642,969	-	-	-	-	-
12	Reports and reproductions	155,062	-	147,906	24	-	2,106
13	Interviewer services	653,515	-	-	1,217	134,931	486,071
14	Interviewer expenses	431,835	-	-	5,027	92,416	332,246
15	Respondent incentives	114,913	-	•	-	-	114,913
16	Clerical labor	370,655	1,848	577	40	1,872	41,697
17	Clerical labor surcharge	191,849	345	247	14	1,076	23,351
18	Miscellaneous	13,892	-	60	118	150	1,640
19	Overtime expenses	37,539	1,923	28	895	1,620	12,596

<sup>1</sup>Legend for project tasks:

'Legend for project tasks:
1 = Survey sampling.
2 = Instrument and materials development.
3 = Field preparations.
4 = Survey training.
5 = Data collection.
6 = Control system development and production.
7 = Data receipt, editing, and document control.

8 = Data coding operations.
9 = Data entry operations.
10 = Control card development, maintenance, and production.
11 = Summary development, maintenance, and production.
12 = Other data processing operations.
13 = Database construction.
14 = Project administration.

Table 15.	Summary of estimated	l costs of project tasks	for linked household design B-Con.
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	Project task <sup>1</sup> —Con.									
	14	13	12	11	10	9	8	7	6	
01	\$223,380	\$382,754	\$584,065	\$157,488	\$99,255	\$357,472	\$67,306	\$271,883	\$205,192	
02 03	189,873 -	155,299 -	291,767 -	57,092 -	38,163	83 -	12,464 -	44,828	76,755 -	
04	33,507	227,455	292,298	100,396	61,092	357,389	54,842	227,055	128,437	
05 06 07	1,675 - 6,701	61 - 804	9,586 1,990 2,885	6,502 21 126	4,921 - 225	233 116,102 17	1,439 61 91	4,975 6,332 54,412	923 16 3	
08	20,104	142	4,449	1,098	839	2	293	1,490	863	
09 10	-	225	21,380	- 819	3,971	-	-	•	1,419	
11 12 13	5,026	226,091 - -	174,394 - 31,296	82,662 - -	46,541 - -	-	-	-	113,281	
14 15	-	-	-	2,146	-	-	-	-	-	
16 17	-	52 28	33,108 11,506	69 34	13 4	129,456 106,230	41,133 11,210	120,753 37,804	37	
18 19	-	52	12 1,692	141 6,778	- 4,578	156 5,193	7 608	110 1,179	11,498 397	

Table 16.	Summary o	of estimated	l costs o	f project	tasks for	linked house	hold design C

		Project task <sup>1</sup>							
	Cost category	Total	1	2	3	4	5		
01	Total	\$7,182,341	\$72,697	\$288,961	\$52,723	\$856,020	\$2,393,797		
	Direct technical labor								
02 03	On-site Off-site	1,574,753 344,587	64,592	21,296	13,164 21,278	37,460 62,579	299,259 260,730		
	Other direct cost								
04	Total	5,263,001	8,105	267,665	18,281	755,981	1,833,808		
05 06 07	Materials and supplies Services Shipping and communications Travel:	79,174 280,409 226,513	241 455 440	1,127 1,134 684	658 1,582 4,971	587 30,541 14,679	32,696 29,004 101,748		
08 09 10	On-site Off-site Consultants	56,813 358,776 27,914	1,043 - -	1,986 147 -	750 6,586 -	14,530 286,792	8,201 65,251 -		
11 12	Computer services Reports and reproductions	986,252 269,012	-	- 261,601	- 12	-	- 2,134		
13 14 15	Interviewer services Interviewer expenses Respondent incentives	1,083,210 687,505 203,245	-	-	608 2,514	239,148 163,797	793,042 517,605 203,245		
16 17	Clerical labor Clerical labor surcharge	612,805 317,962	2,914 544	629 263	20 7	2,326 1,356	42,283 23,689		
18 19	Miscellaneous	21,642 51,769	۔ 2,468	64 30	67 506	187 2,038	1,762 13,148		

<sup>1</sup>Legend for project tasks: 1 = Survey sampling. 2 = Instrument and materials development.

3 = Field preparations. 4 = Survey training.

5 = Data collection.
6 = Control system development and production.
7 = Data receipt, editing, and document control.

8 = Data coding operations.
9 = Data entry operations.
10 = Control card development, maintenance, and production.
11 = Summary development, maintenance, and production.
12 = Other data processing operations.
13 = Database construction.
14 = Project administration.

Table 16. Summary of estimated costs of project tasks for linked household design C---Con.

	Project task <sup>1</sup> —Con.										
	14	13	12	11	10	9	8	7	6		
01	\$234,018	\$552,884	\$847,568	\$246,811	\$143,776	\$632,080	\$112,686	\$437,744	\$310,576		
02 03	\$198,915 -	224,311	387,998 -	86,349 -	55,263 -	143	16,564 -	54,234 -	115,205		
04	35,103	328,573	459,570	160,462	88,513	631,937	96,122	383,510	195,371		
05 06 07	1,755 - 7,021	88 - 1,162	15,473 3,267 2,885	9,631 35 126	7,131 - 225	407 205,347 30	1,913 89 161	6,277 8,929 92,378	1,190 26 3		
08		206				30			-		
09	21,062	200	4,449	1,098	839	3	293	1,490	863		
10	-	325	21,380	819	3,971	-	-	-	1,419		
11	-	326,602	284,804	113,497	68,955	-		-	172,394		
12	5,265		•	-	•	-	-	-	•		
13	-	-	50,412	-	-	-	-	-	-		
14	-	-	-	3,589	-	-	-	-	-		
15	-	-	-	•	-	-	•	-	•		
16	-	75	55,034	88	17	228,970	72,751	207,650	48		
17	-	40	19,139	46	5	187,891	19,827	65,155	-		
18 19	-	- 75	20 2,707	204 11,329	- 7,370	272 9,017	12 1,076	135 1,496	18,919 509		

Table 17.	Summary of estimated costs of	f project tasks for linked household design D
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				Proje	ct task <sup>1</sup>						
	Cost category	Total	1	2	3	4	5				
01	Total	\$6,962,291	\$69,003	\$266,142	\$93,672	\$792,966	\$2,445,596				
	Direct technical labor										
02 03	On-site	1,568,976 395,464	61,373 -	21,034	22,397 37,922	35,954 60,050	339,466 297,492				
	Other direct cost										
04	Total	4,997,851	7,630	245,108	33,353	696,962	1,808,638				
05 06 07	Materials and supplies Services Shipping and communications Travel:	79,859 263,659 234,648	224 428 412	1,111 1,122 676	1,152 3,163 8,742	563 29,288 13,732	35,432 30,242 113,469				
08 09 10	On-site	61,649 350,140	998 -	1,962 145	1,302 11,659	13,946 265,307	13,151 73,029				
10	Consultants Computer services	27,894 918,373	-	-	-	-	-				
12 13	Reports and reproductions	247,043 1,019,956	-	239,122	24 1,217	218,676	2,632 753,431				
14 15	Interviewer expenses	661,210 185,781	-	-	5,027	149,776	503,104 185,781				
16 17	Clerical labor Clerical labor surcharge	576,768 299,140	2,703 505	618 260	40 14	2,237 1,301	52,117 29,191				
17 18 19	Miscellaneous	299,140 20,394 51,337	2,360	63 29	14 118 895 `	1,301 180 1,956	1,977 15,082				

<sup>1</sup>Legend for project tasks:
1 = Survey sampling.
2 = Instrument and materials development.
3 = Field preparations.

4 = Survey training.

5 = Data collection.

6 = Control system development and production.
7 = Data receipt, editing, and document control.

8 = Data coding operations.

9 = Data entry operations. 10 = Control card development, maintenance, and production.

11 = Summary development, maintenance, and production.
12 = Other data processing operations.
13 = Database construction.
14 = Project administration.

Table 17. Summary of estimated costs of project tasks for linked household design D-Con.

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				n.	oject task <sup>1</sup> —Col	Pi			
	14	13	12	11	10	9	8	7	6
01	\$234,018	\$519,241	\$795,467	\$229,146	\$134,972	\$577,786	\$103,712	\$410,830	\$289,740
02 03	198,915	210,667	368,971 -	80,563 -	51,881 -	131	15,753 -	54,267 -	107,604 -
04	35,103	308,574	426,496	148,583	83,091	577,655	87,959	356,563	182,136
05 06	1,755	83	14,309 3,014	9,013 32	6,694	373 187,704	1,819 83	6,194 8,559	1,137 24
07	7,021	1,091	2,885	126	225	27	147	86,092	3
08	21,062	193	4,449	1,098	839	3	293	1,490	863
09	-	•	-	-	-	-	-	-	-
10	-	305	21,380	819	3,971	-	-	-	1,419
11		306,724	262,975	123,444	64,524	-	-	-	160,706
12	5,265	-		-	-	-	-	-	-
13	-	-	46,632		-	-	-	-	-
14	-	-	•	3,303	-	-	-	•	-
15	-	- 70	50,699	84	- 16	- 209,294	- 66,500	192,344	- 46
16 17	-	38	17,629	43	4	171,744	18,123	60,288	40
18	-		17,029	192	4	249	10,123	135	17,451
19	-	70	2,506	10,429	6,818	8,261	983	1,461	487

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# Chapter 5 An optimally allocated design

The designs previously described are self-weighting and selected by aggregating the National Health Interview Survey (NHIS) sample over a short time period. Cost savings result from linking these designs to NHIS, but they are not substantial. One reason for the lack of substantial cost savings is that these designs include little of the available NHIS information. Using the characteristics of NHIS respondents, greater savings are possible by stratification and optimal allocation of the sample.

To investigate this, five optimally allocated linked household designs were studied. Two designs are optimally allocated self-weighting designs, one with the precision of the 6,000 originating base reporting unit (OBRU) unlinked design, the other with the precision of the 10,000-OBRU design. Next, the selfweighting constraint was removed for two optimally allocated designs, one using the 6,000-OBRU constraints, the second using the 10,000-OBRU constraints. Because increasing the sample size to 10,000 OBRU's improves precision for smaller domains such as medicaid recipients, a fifth design was developed using the 6,000-OBRU constraints for the total population and the 10,000-OBRU constraints for the medicaid subpopulation.

#### Definition

Stratification of the sample is usually proportional to stratum size, except when oversampling of certain population subgroups is specified. However, because data collection costs and variances differ among strata, optimal allocation of the sample may result in substantial cost savings. For the National Medical Expenditure Survey (NMES), a multipurpose survey with many outcome measures and reporting domains, the preferred optimization strategy is one that minimizes total survey cost subject to multiple variance constraints. Separate variance constraints are set to control the precision of key survey statistics for the total population and for important reporting domains.

To optimally allocate the sample among strata, cost and variance models are needed. The following linear function is used to model survey costs for a sample design with L sample size levels, m(l):

$$C = C_0 + \sum_{l=1}^{L} C(l)m(l)$$
(14)

where C =total survey cost

 $C_0 =$  fixed administrative cost of the survey

- C(l) = cost of surveying a unit from the lth design level where l may index a combination of design stages, phases, and strata
- m(l) = sample size for the *l*th design level

The corresponding variance model for a particular statistic and domain k is

$$V_{k} = \sum_{l=1}^{L} \frac{V_{k}(l)}{m(l)}$$
(15)

where  $V_k$  = variance of the domain k statistic

 $V_k(l)$  = variance component associated with the kth domain and sampling from the *l*th design level

These cost and variance models illustrate that as the sample size for each stratum increases, the variance decreases as the total cost of the survey increases.

To determine the optimum sample sizes for the L design levels, the maximum variances  $(V_k^*)$  allowed for the designated domain k estimates must be specified. This may be represented mathematically as the set of level-specific sample sizes m(l)that minimize the total survey cost C subject to  $V_k \leq V_k^*$  and  $m(l) \geq 0$  for all l. For a single variance constraint problem, the optimal allocation to level l is

$$m(l) = \left[\frac{V(l)}{C(l)}\right]^{1/2} \sum_{l=1}^{L} \frac{[V(l)/C(l)]^{1/2}}{V^*}$$
(16)

With optimum allocation, these level-specific sampling rates tend to increase as the associated variance increases or the data collection cost decreases.

Few surveys are conducted to obtain a single estimate. For sample allocation based on the single variance constraint solution, several estimates would be considered and the design would be optimized for only one. The preferred strategy simultaneously considers several estimates chosen by classifying the survey statistics according to their variance properties and selecting a typical variance model from each class. Unlike the single constraint case, optimization for multiple variance constraints does not have a closed form solution; Cochran  $(pp. 119-123)^3$  reviews a number of approaches to obtain solutions for these problems.

The NMES optimization was obtained using an optimization approach developed by Chromy, described in reference 4. Chromy's optimization algorithm is an iterative approach that provides an optimal solution when the convergence criteria are met.

NHIS household sampling units provide useful information for NMES. This information is generally person-level such as age, race, sex, relationship to head of household, limitation of activity, bed disability days, perceived health status, medical conditions, education level, marital status, and employment. Because NMES samples entire households to facilitate familylevel analysis, these data must be aggregated to the household level for stratification.

Stratification of the NHIS sample before selection of the NMES sample provides control over the distribution of the sample while increasing the precision of survey estimates. The variance of estimates is reduced and the precision increased by sampling stratified to maximize the between-stratum variation and minimize the within-stratum variation. Variables used for stratification should result in homogeneity of the units within strata and heterogeneity between strata.

Time constraints prevented the examination of 1980 National Medical Care Utilization and Expenditure Survey (NMCUES) data to determine which variables should be used for stratification of the NHIS sample before NMES sample selection. Instead, variables that are considered good predictors of health care utilization and expenditures were used for stratification. These variables are black and all other races, aged and not aged, poor and not poor, and self-perceived health status (healthy and not healthy). Sample size limitations of the 1980 NMCUES database used to estimate variance components required collapsing of the black strata over the poverty variable, resulting in eight all-other-race strata and four black strata.

To demonstrate the advantages of an optimum allocation approach, five optimal designs were developed. The domains that were included in the optimization are the total population and medicaid recipients. For use in stratification, dichotomous OBRU-level variables denoted race (black versus nonblack), poverty status (more or less than 150 percent of the official poverty index), age status (containing no person greater than or equal to 65 years versus containing at least one), and health status (containing no person with poor or fair health versus containing at least one). The optimization was conducted for nine utilization and expenditure rates and for the subpopulation with large out-of-pocket expenses. First, variance modeling for a stratified, linked household design drawn from the first phase NHIS sample was conducted. Second, the cost component for each second phase stratum and each stage of the first phase NHIS design was modeled. Finally, optimization was conducted and its results assessed. The optimization program computes the total survey costs for the optimal design based on the unit costs. Because the total cost was available, full scale costing to evaluate the design was not necessary. Therefore, this step was eliminated for all the optimally allocated designs.

#### Variance modeling

Using a stratified sampling approach, NMES would estimate the mean for domain k as

$$\widetilde{Y}_{k}(\text{NMES}) = \sum_{h=1}^{H} \hat{\pi}_{k}(h) \overline{Y}_{k}(h)$$
 (17)

where  $\overline{Y}_{k}(h) = \text{NMES}$  estimated mean for stratum h

- $\hat{\pi}_k(h) =$ NHIS-estimated fraction of the kth subpopulation total person-years associated with the hth stratum
  - H = number of sample strata

For the nine utilization and expenditure measures, the stratum mean is estimated as

$$\overline{Y}_{k}(h) = \frac{\sum_{i \in h} W(i)\delta_{k}(i)Y(i)}{\sum_{i \in h} W(i)\delta_{k}(i)T(i)}$$
(18)

where W(i) = sampling weight of the *i*th person

- $\delta_k(i) = 1$  if the *i*th person belongs to the *k*th domain and 0 if not
- Y(i) = response of the *i*th person
- T(i) = fraction of the year that the *i*th person was eligible for NMES

For the proportion burdened with large out-of-pocket expenses, the stratum mean is estimated as

$$\overline{Y}_{k}(h) = \frac{\sum_{i \in h} W(i)\delta_{k}(i)T(i)Y(i)}{\sum_{i \in h} W(i)\delta_{k}(i)T(i)}$$
(19)

where Y(i) = 1 if the annualized out-of-pocket expenses are large (more than \$200) and 0 if not.

To simplify modeling the variance, it is assumed that NHIS oversampling of black persons is at the last stage and that black and all other races is a stratification variable. Therefore, the variance of the stratified estimate is modeled as

$$Var [\overline{Y}_{k}(NMES)] = Var_{NHIS} \{E[\overline{Y}_{k}(NHIS)]\} + E_{NHIS} \{Var [\overline{Y}_{k}(NHIS)]\} = Var_{NHIS} [\overline{Y}_{k}(NHIS)] + E_{NHIS} \sum_{k=1}^{H} \frac{\pi_{k}^{2}(h)S_{k}^{2}(h)[1-f(h)]}{m(h)}$$

 $\hat{}= D_w(k)$ 

Because

$$\theta_B = \frac{1.4\pi_B}{1.4\pi_B + \pi_{AOR}} \tag{24}$$

and

$$\theta_{AOR} = \frac{\pi_{AOR}}{1.4\pi_B + \pi_{AOR}}$$
(25)

 $D_{w}(k)$  may also be expressed as

$$D_{w}(k) = 1 + \frac{0.16\pi_{B}\pi_{AOR}}{1.4}$$
(26)

For convenience, relative variance components are used in the optimization. To model the relative variances,

$$RV_{k}(\text{NMES}) = \frac{\text{Var}\left[Y_{k}(\text{NMES})\right]}{\overline{Y}_{k}^{2}(\text{NMES})}$$
(27)

For domain k, the relative variance of a mean estimated using the linked household design can be expressed as

$$RV_{k}(\text{NMES}) = \sum_{l=1}^{H} \frac{RV_{k}(l)}{m(l)} + \sum_{l=H+1}^{H+2} \frac{RV_{k}(l)}{m(l)}$$
(28)

where l = 1, 2, ..., H are the second phase strata used in selecting the NMES subsample, and H + 1 and H + 2 are the first phase segment and PSU sampling stages.

#### **Cost modeling**

If C(l) represents the variable unit cost for a selection from level l, then the optimization problem may be stated as follows: Minimize

$$CV(NMES) = \sum_{l=1}^{H+2} m(l)C(l)$$
 (29)

subject to

1. 
$$\sum_{l=1}^{H+2} \frac{RV_k(l)}{m(l)} \le RV_k^* \text{ for } k = 1, 2, \dots, K$$

- 2.  $m(l) \ge 0$  for l = 1, 2, ..., H+2
- 3.  $200 \le m(H+2) \le m(H+1)$
- 4.  $m(l) \le m(H+1)$  for l = 1, 2, ..., H

where 
$$CV(NMES) =$$
 total variable cost for NMES

 $RV_k^* =$  relative variance constraint for the kth domain

$$\cdot \left[ \frac{\sigma_k^2(\text{PSU})}{r} + \frac{\sigma_k^2(\text{SEG})}{r\bar{s}} + \frac{\sigma_k^2(\text{OBRU})}{r\bar{s}\bar{t}} \right]^{H}$$

$$+\sum_{h=1}^{\frac{\pi_{k}(h)S_{k}^{*}(h)[1-f(h)]}{E[m(h)]}}$$
(20)

- where  $D_w(k) =$  design effect for NHIS unequal weighting for the kth domain
  - $\sigma_k^2(\text{PSU}) =$  between NHIS primary sampling unit (PSU) variance component for domain k
  - $\sigma_k^2$ (SEG) = between NHIS segment, within NHIS PSU variance component for domain k
  - $\sigma_k^2(\text{OBRU}) = \text{between NHIS OBRU}$ , within NHIS segment variance component for domain k
    - $S_k^2(h) =$  stratum h variance for domain k
    - f(h) =NMES subsampling rate for stratum h or m(h)/n(h)
    - m(h) = NMES stratum h OBRU sample size
    - n(h) = NHIS stratum h OBRU sample size

The variance components computed from the 1980 NMCUES were used to estimate the NHIS components. A Taylor series approximation for the simple random sampling variance of a combined ratio estimator was used to estimate  $S_L^2(h)$ .

The expected NMES sample size from the hth stratum can be expressed as

$$E[m(h)] = r\bar{s}\bar{t}f(h)\pi'(h)$$
(21)

where  $\pi'(h) =$  expected fraction of the NHIS sample from the *h*th strata or

$$\pi'(h) = \frac{M(h)o(h)}{\sum_{h=1}^{H} M(h)o(h)}$$
(22)

and M(h) is the population count of OBRU's in stratum h.

Assuming that black and all other races is used as a stratification variable with equal probability sampling within strata, the design effect for unequal weighting in domain k estimation is modeled as

$$D_{w}(k) = \frac{\pi_{B}^{2}}{\theta_{B}} + \frac{\pi_{AOR}^{2}}{\theta_{AOR}}$$
(23)

where  $\pi_B =$  proportion of black persons in the population

 $\pi_{AOR}$  = proportion of all other races in the population

 $\theta_B$  = proportion of black persons in the NHIS sample  $\theta_{AOB}$  = proportion of all other races in the NHIS sample The variable costs for the PSU stage of sampling [C(H + 2)]and the segment level of sampling [C(H + 1)] were obtained by aggregating the task-level unit costs determined by the cost modeling of the self-weighting linked household design cost modeling (chapter 4). The unit costs for the subsampled OBRU's within NHIS-defined strata vary depending on the response and movement rates within the strata. In a procedure similar to that described in chapter 4 for the total population, the 1980 NMCUES experience was used to estimate the rates at which ineligibles, nonrespondents, and movers are encountered and to develop the OBRU-level cost component for each of the 12 strata. The unit costs developed for the selfweighting linked household design for tracing movers, interviewing ineligibles, and interviewing outside and inside the clusters were used in forming the total unit costs for each stratum.

#### **Optimization results**

The first design investigated is a stratified, self-weighting linked household design. Using this design, the variance is expressed as in equation (20) where f(h) = f/o(h). The factor f is the subsampling rate desired for the NMES subsample of NHIS after NHIS oversampling is removed. The Chromy optimization procedure was used to obtain optimum values for the number of PSU's, the average number of segments to sample per PSU, and the NMES subsampling rate used within the sample segments  $(r, \bar{s}, \text{ and } f)$ . For use in the optimization, the simplified variance function is recast in the form of equation (15) as

$$\operatorname{Var}\left[\overline{Y}_{k}(\operatorname{NMES})\right] = \frac{\left[D_{w}(k)\sigma_{k}^{2}(\operatorname{PSU})\right]}{r} + \frac{D_{w}(k)\sigma_{k}^{2}(\operatorname{SEG}) + D_{w}(k)\sigma_{k}^{2}(\operatorname{OBRU})/\overline{t}}{r\overline{s}} - \frac{\sum_{h=1}^{H}\pi_{k}^{2}(h)S_{k}^{2}(h)/\pi'(h)\overline{t}}{r\overline{s}} + \frac{\sum_{h=1}^{H}\pi_{k}^{2}(h)S_{k}^{2}(h)o(h)/\overline{t}\pi'(h)}{r\overline{s}f}$$
(30)

Correspondingly recasting the linear cost model leads to H second phase stratum cost parameters of the form

$$C'(l) = \frac{C(l)t\pi'(l)}{o(l)}$$
(31)

The optimization was performed twice. When the variance constraints associated with the 6,000-OBRU unlinked design were used, the optimal solution was 102 PSU's, 1,258 segments, and 5,980 responding OBRU's. With a subsampling rate f of 83 percent, black strata are subsampled at a 59-percent rate

(f/1.4) and all-other-race strata at the 83-percent rate. The total cost for the design is \$4,844,013 compared with \$4,963,013 for the unlinked design with the same precision.

When the variance constraints associated with the 10,000-OBRU unlinked design are used, the optimal stratified linked household design has 103 PSU's, 2,117 segments, 9,960 responding OBRU's, and a subsampling rate f of 82 percent. Allowing for the NMES oversampling, black strata are subsampled at a 58-percent rate and all-other-race strata at the 82percent rate. The total cost for this design is \$6,931,233 compared with \$7,209,409 for the unlinked design with the same precision.

The stratified household design, with 10,000-OBRU precision, incorporates 103 PSU's. The unstratified design, previously described in chapter 3, is most cost efficient with 200 PSU's. This difference is the result of instability of the estimated variance components used to obtain the sample sizes for the unstratified designs.

The next set of designs investigated are the stratified linked household designs without the self-weighting constraint. The advantage of this type of design is that heavy utilizers of health care services can be identified and oversampled. For use in the optimization, the variance given in equation (20) was recast following equation (15).

$$\operatorname{Var}\left[\overline{Y}_{k}(\operatorname{NMES})\right] = \frac{D_{w}(k)\sigma_{k}^{2}(\operatorname{PSU})}{r} + \frac{D_{w}(k)\sigma_{k}^{2}(\operatorname{SEG}) + D_{w}(k)\sigma_{k}^{2}(\operatorname{OBRU})/\overline{t}}{r\overline{s}} - \frac{\sum_{h=1}^{H}\pi_{k}^{2}(h)S_{k}^{2}(h)/\pi'(h)\overline{t}}{r\overline{s}} + \sum_{h=1}^{H}\frac{\pi_{k}^{2}(h)S_{k}^{2}(h)/t\pi'(h)}{r\overline{s}f(h)}$$
(32)

To optimize over PSU's (r), segments  $(r\overline{s})$ , and NMES strata (h = 1, 2, ..., H), the stratified linked sample has H + 2 design levels. Using expression (32) for the variance, revised unit costs are computed for each of the H second phase strata or

$$C'(l) = C(l)t\pi'(l) \tag{33}$$

The total population and medicaid recipients are used in the optimization. Medicare recipients, the poor, and those in families with college educated heads of households were not included because an instability of the variance components was observed with negative segment-level variance components for some domain estimates. Due to time constraints, examination and correction of the negative components were not possible.

First, an optimally allocated design with the precision constraints of the unlinked 6,000-OBRU design for the total and medicaid domains was investigated. The optimal solution used 98 PSU's, 1,152 segments, and 5,880 responding OBRU's with subsampling rates ranging from 57–100 percent. In general, the not healthy and all-other-race groups are sampled at a higher rate than is the black group. Greater percents of NHIS all other race persons are selected than black persons because the number of black persons occurs at a rate 1.4 times greater than that for persons of all other races in the NHIS sample. The total cost for this design is \$4,770,353 compared with \$4,963,013 for the unlinked 6,000-OBRU design and \$4,844,013 for the self-weighting optimally allocated design.

Next, an optimally allocated design with the precision of the 10,000-OBRU unlinked design for the total and the medicaid domains was investigated. The optimal solution used 106 PSU's, 1,811 segments, and 9,717 responding OBRU's with subsampling rates ranging from 59–100 percent. The total cost for the design is \$6,758,063 compared with \$7,209,409 for the 10,000-OBRU unlinked design and \$6,931,233 for the optimally allocated self-weighting design.

For household samples drawn from area frames, there is little information available for use in sample stratification. To obtain the required sample sizes for small domains, a sample size larger than usual is frequently used. With household-level stratification information, these small domains can be oversampled without increasing the size of the total sample.

To illustrate this advantage, an optimally allocated design, with the precision of the unlinked 10,000-OBRU design for the medicaid domain and of the 6,000-OBRU design for total population estimates, was developed. These constraints result in an optimal design with 95 PSU's, 2,092 segments, and 7,228 responding OBRU's with NMES subsampling rates ranging from 32-100 percent. The total cost for the design with 6,000 and 10,000 OBRU's is \$5,601,533, which compares well with the \$6,758,063 cost for the comparable not-self-weighting design with 10,000-OBRU constraints for both the total and medicaid domain statistics. Tables 18–20 summarize the results of these comparisons.

#### Other design considerations

NMES will have many small analysis domains including the medicaid, the medicare, the aged, the poor, and the black populations. In the past, separate analyses have been made possible by selecting self-weighting samples large enough to obtain adequate precision for these domains. This approach results in precision greater than necessary for large domains such as the not-aged or white domains. Without linkage, however, this is the best approach because household characteristics are not available for use in sampling.

Although beyond the scope of this report, precision constraints for the NMES should be set for a large group of policyrelevant domains. With linkage to NHIS, there is much information about households that can be used to create an optimally allocated design with increased precision for selected domains. The stability of the variance components and the accuracy of the cost components should also be considered. Finally, cost modeling should include the effect of the aggregation length of the NHIS sample. The reporting domains to be included in the optimization need careful attention. Precision is assured for statistics and domains included in the optimization. The precision for other statistics and domains will depend on their relation to the statistics and domains included in the optimization.

The optimizations were designed for total utilization and expenditure statistics for the total population and for the medicaid population. The stratified self-weighting linked household design insures precision for these domains and statistics by selecting a self-weighting design with a sufficient sample size. In the stratified linked household designs without the selfweighting constraint, the precision for these statistics was maintained and the cost decreased by oversampling the poor and the not healthy and undersampling the not poor and the healthy. For domains and statistics not included in the optimization, neither of these optimal designs may yield statistics of the desired precision.

Examples from the optimizations described in this chapter demonstrate this point. The stratified self-weighting linked design, optimized for the variance constraints of the 6,000-OBRU unlinked design, may not produce estimates of the desired precision for small domains such as newborns. Using the variance constraints for the 10,000-OBRU unlinked design, the sample size for newborns still may not be sufficient to support detailed analyses. Increasing the sample size of the self-weighting design yields increased precision for such small domains and greater precision than necessary for large domains.

Without the self-weighting constraint, an optimally allocated design can be created that obtains the desired precision for a small domain by oversampling from strata where domain members are concentrated. If the 10,000-OBRU unlinked design yields the required variance constraints for the medicaid domain, the self-weighting linked design to use is that which vields the variance constraints of the 10.000-OBRU unlinked design for all domains. If the 6,000-OBRU unlinked design vields variance constraints acceptable for the total population. the not-self-weighting optimally allocated linked design can achieve both sets of variance constraints by oversampling strata with a high concentration of medicaid recipients. The survey costs with the not-self-weighting approach (the notself-weighting design with 6,000 total and 10,000 medicaid precision constraints in table 19 are \$5,601,533 compared with \$6,931,233 for the self-weighting design (the self-weighting design with 10,000 and 10,000 respondents in table 19).

The disadvantage of the optimally allocated not-selfweighting approach is that it may not yield estimates of the desired precision for domains and statistics not included in the optimization. The not-self-weighting design with 6,000 total and 10,000 medicaid precision constraints produces estimates of the desired precision for the total utilization and total expenditure statistics by oversampling from the not healthy strata. If total income is being estimated instead, estimates of the desired precision can not be assured because the design does not control for the precision of income estimates. Alternatively, if total utilization or total expenditures are being estimated for a domain not included in the optimization, such as the medicare domain, the design may not yield estimates of the desired precision. The precision of estimates for domains and statistics not included in the optimization depends on their relation to the statistics and domains included in the optimization.

Although most surveys include many domains and statistics, this does not preclude use of a not-self-weighting optimally allocated design. A strategy using this design is to consider several estimates chosen by classifying their variance properties and selecting a typical variance model from each class. Similarly, the domains to include in the optimization can be chosen by listing the important domains and selecting those that represent diverse groups of the population.

Because extreme groups are usually rare, they must be represented in the set of domains subject to optimization to obtain an adequate sample size. For example, a survey comparing health expenditures for different income groups should include the poor and the wealthy as domains in the optimization. It may not be necessary to include the large middle income portion of the population as a domain, particularly if the total population is included as a domain in the optimization.

Linkage of NMES to NHIS makes available the names, addresses, and personal characteristics of sample households before data collection. The design with the most potential for using this information is the stratified not-self-weighting optimally allocated design. Research to produce this design would determine the domains and statistics of interest to the survey and the appropriate set to include in the optimization. The 1980 NMCUES data could be used in constructing variance and cost models. The advantages of implementing an optimally allocated design should far exceed the costs of its development.

#### Table 18. Sample sizes for the alternate optimally allocated designs

Design type	Primary sampling units	Segments	Originating base reporting units (OBRU's)	Cost
Self-weighting, precision of 6,000 OBRU design	102	1.258	5,980	\$4,844,013
Self-weighting, precision of 10,000 OBRU design	103	2,117	9,960	6,931,233
Not-self-weighting, precision of 6,000 OBRU design	98	1,152	5,880	4,770,353
Not-self-weighting, precision of 10,000 OBRU design	106	1,811	9,717	6,758,063
Not-self-weighting, precision of 6,000 total OBRU and 10,000 medicaid OBRU designs	95	2,092	7,228	5,601,533

#### Table 19. Stratum sampling rates for the alternate optimally allocated designs

			Design type		
Strata	Self-weighting, precison of 6,000 OBRU design	Self-weighting, precision of 10,000 OBRU design	Not self-weighting, precision of 6,000 OBRU design	Not self-weighting, precision of 10,000 OBRU design	Not self-weighting, precision of 6,000 total OBRU and 10,000 medicaid OBRU designs
All other races, not aged, not poor, healthy	83	82	86	94	41
All other races, not aged, not poor, not healthy	83	82	99	99	95
All other races, not aged, poor, healthy	83	82	76	79	63
All other races, not aged, poor, not healthy,	83	82	100	100	100
All other races, aged, not poor, healthy	83	82	83	84	32
All other races, aged, not poor, not healthy	83	82	100	100	93
All other races, aged, poor, healthy,	83	82	88	97	72
All other races, aged, poor, not healthy	83	82	87	77	72
Black, not aged, healthy	59	58	61	67	59
Black, not aged, not healthy	59	58	76	79	76
Black, aged, healthy	59	58	57	59	35
Black, aged, not healthy	59	58	100	100	100

OBRU = originating base reporting unit.

#### Table 20. Stratum originating base reporting unit (OBRU) sample sizes for the alternate optimally allocated designs

			Design type		
Strata	Self-weighting, precison of 6,000 OBRU design	Self-weighting, precision of 10,000 OBRU design	Not self-weighting, precision of 6,000 OBRU design	Not self-weighting, precision of 10,000 OBRU design	Not self-weighting, precision of 6,000 total OBRU and 10,000 medicaid OBRU designs
All strata	5,980	9,960	5,880	9,717	7,228
All other races, not aged, not poor, healthy	2,826	4,707	2,697	4,622	2,328
All other races, not aged, not poor, not healthy	556	927	612	957	1,069
All other races, not aged, poor, healthy	451	751	380	625	574
All other races, not aged, poor, not healthy	250	416	277	435	503
All other races, aged, not poor, healthy	461	768	422	674	298
All other races, aged, not poor, not healthy	279	464	309	486	524
All other races, aged, poor, healthy	268	446	262	454	390
All other races, aged, poor, not healthy	265	441	256	356	383
Black, not aged, healthy	351	585	332	573	586
Black, not aged, not healthy	152	254	179	293	328
Black, aged, healthy	52	87	47	75	52
Black, aged, not healthy	69	114	107	167	193

OBRU = originating base reporting unit.

## Chapter 6 Comparison of the designs and recommendations

The National Medical Expenditure Survey (NMES) design types investigated in this study have similar features. Regardless of how the sample is selected, all of the designs assume that each sample household is interviewed personally in rounds 1, 2, and 5, and that the telephone is used whenever possible in rounds 3 and 4.

Each design defines key persons to be followed for all rounds of data collection. The designs also define key persons as those who, in rounds 2–5, are either born or return from the military, overseas, or a long-term care institution and enter an existing family. All other persons who are members of families formed by key persons are classified as nonkey. Nonkey persons have data collected for them only as long as they belong to families with members who are key persons. The data for key persons are used for person-level analyses; nonkey person data are only used to construct aggregated data used in family-level analyses.

In round 1, a household roster is obtained, and health care data are collected for all household members including college students living away from home. During the first interview, the household is given a calendar diary and instructed as to its use. An incentive of \$5 is paid to the household and its members are advised that another \$5 will be paid to them at the end of the survey. The household is advised that a summary of the reported health care data will be mailed to its members before each interview so that erroneous or missing information can be corrected.

Round 2 is also conducted by personal interview for the design types investigated in this study. The advantages of a second personal interview round are that the interviewer can review the summary with the respondent; and, because the bulk of survey attrition occurs at round 2, a personal interview should reduce the level of attrition early in the survey and commit the respondent to the survey.

The next two rounds of data collection use the telephone whenever possible. Because round 4 is at the end of the year, not all respondents are included. Because December 31 is the end of the survey reference period, approximately 30 percent of the sample is not interviewed in round 4 but, instead, early in round 5 (that is, shortly after January 1 of the next year).

The fifth and final round of data collection is conducted by personal interview. In addition to obtaining the health care data through December 31 of the past year, the round 5 interview obtains annual income and other data that are not available until after the end of the reference period.

The same target population definition is used by the Na-

tional Health Interview Survey (NHIS) and NMES, which facilitates using the NHIS sample as a frame for NMES. Both surveys define their target populations as the civilian noninstitutionalized residents of the United States. NHIS is based on a national area sample of housing units and group quarters and is similar to the 1980 National Medical Care Utilization and Expenditure Survey (NMCUES) design except for the sampling of college students. NHIS includes college students in the sample when their college residence is sampled. Because of its interest in family-level analyses, NMES links college students who are single, 17-22 years of age, and living away from home to their parents' residence. Only when the parents' residence is selected is the college student included in the sample. The difference between the definitions does not present problems for linkage of NMES to NHIS provided that NHIS identifies all college students who are single, 17-22 years of age, and living away from home and asks sample NHIS families to provide name and address information for these college students.

Four types of sample designs were investigated in this study, including two unlinked designs, four linked NHIS and NMES dwelling unit designs, four linked NHIS and NMES household designs, and five optimally allocated linked household designs. Table 21 summarizes the sample size and cost for the 1980 NMCUES and for the 14 designs investigated for use in the 1987 NMES. The cost of the five optimally allocated designs compares well with that of the other designs. These costs were constructed from the 1980 NMCUES experience and are not adjusted for inflation.

Table 21 includes the months that the NHIS sample must be aggregated to obtain the required number of sample segments from the specified number of primary sampling units (PSU's). These estimates of aggregation time are based on the assumptions that NHIS includes 8,750 segments and 200 PSU's for an average of 43.75 segments per PSU in a year and that NMES is selected from the 90 percent consisting of personal interviews. The aggregation times range from 1.5-6.7 months; the longer periods of aggregation are required for the optimally allocated designs. Modeling of movement is only approximate, so the costs associated with movement may be understated, particularly for designs that aggregate over a longer period of time. More attention needs to be given to cost modeling of movement as the time between NHIS and NMES increases.

In modeling the costs for the designs it is assumed that the NMES contractor selects the sample. The NHIS interviewer in the NMES-subsampled segments is given a three-part tearsheet on which to record the information needed in the NMES sample selection. This information includes names and addresses, NHIS-identifiers, and person characteristics needed for stratification. The tearsheet is completed at the time of NHIS data collection. The tearsheets are distributed on a flow basis, one copy to the contractor, one copy to the U.S. Bureau of the Census field office, and one copy to the interviewer's records. With this approach, the contractor constructs the frame on a flow basis. The Census field office also reviews the documents on a flow basis and advises the contractor of any discrepancies. With the tearsheet approach, the NMES sample can also be selected by the U.S. Bureau of the Census or the National Center for Health Statistics.

For costing the sampling effort, it is assumed that the contractor does the frame construction and sampling. An advantage of selection by the contractor is quality control. NMES is a complex study that requires integration of the effort of sampling statisticians, survey operations specialists, and computer programmers. To coordinate NMES activities and ensure the quality of the product, the contractor should have direct control over all project activities.

The cost savings demonstrated by the optimally allocated designs, particularly the not-self-weighting designs, indicate that there are significant savings possible with NHIS linkage. Further study would be needed to construct such a design for NMES. It is recommended that a full scale design study be conducted before the 1987 NMES to determine the sample size parameters of the design. This study should identify potential high expenditure respondents from NHIS data and use this information to improve the precision of survey estimates to reduce the data collection costs for the survey.

Proposed NMES design parameters should be tested in a pilot study before implementation. This pilot study should test linkage methods, data collection alternatives, and questionnaire changes since the 1980 NMCUES. The use of NHIS-derived information should be considered as a means to reduce the data collection costs of NMES. In this investigation the data collection pattern of the 1980 NMCUES was followed. However, this approach may not be necessary when an NHIS-based list frame is available.

It appears possible that one or more of the personal interview rounds could be replaced by a telephone interview round without adversely affecting response rates. The first round should use personal interviews whenever possible. Personal contact is necessary to establish the creditability of the study, to persuade the respondent to participate, and to instruct the respondent in the use of the calendar diary and the summary. Telephone numbers available from NHIS may be used to make appointments, reducing data collection costs. Before implementing this, the procedure should be tested in a pilot study to determine its impact on response.

Another strategy that could be tested is using NHIS to obtain round 1 data for NMES. Using this approach, NHIS families to be included in NMES would have the NHIS instrument administered along with a supplement to obtain the required NMES round 1 data not normally obtained by NHIS. For example, NHIS obtains health care expenditures and utilization data for the week before data collection. The NMES supplement would collect additional data for the period since January 1. If this combined NHIS and NMES interview approach were effective, one round of data collection could be eliminated. If this strategy is considered for NMES, a pilot study should be conducted to determine whether adding a NMES supplement to selected NHIS family interviews would contaminate either NHIS or NMES data. This question of NHIS contamination could be tested by comparing NHIS data collected in the usual manner with NHIS data collected when a NMES supplement was used. The question of the effect on NMES could be tested by comparing NMES data obtained by NHIS interviewers with an NMES supplement with NMES data obtained in an independent NMES interview.

### Table 21. Sample size summary for the alternate National Medical Expenditure Survey (NMES) design

		Sample si			
Design	Primary sampling units	Segments	Originating base reporting units (OBRU's)	Aggregation time	Direct cost
Unlinked designs					
6,000-respondent OBRU's	102 102	750 1,250	6,000 10,000	••••	\$4,963,013 7,209,409
Linked dwelling unit designs					
Design A (6,000-respondent OBRU's) Design B (6,000-respondent OBRU's) Design C (10,000-respondent OBRU's) Design D (10,000-respondent OBRU's)	100 200 100 200	976 921 1,629 1,489	5,856 5,526 9,774 8,934	3.0 1.4 5.0 2.3	4,871,106 4,947,848 7,147,752 6,930,673
Linked household designs					
Design A (6,000-respondent OBRU's) Design B (6,000-respondent OBRU's) Design C (10,000-respondent OBRU's) Design D (10,000-respondent OBRU's)	100 200 100 200	976 921 1,629 1,489	5,856 5,526 9,774 8,934	3.0 1.4 5.0 2.3	4,891,831 4,967,406 7,182,341 7,209,409
Linked stratified optimally allocated household designs					
Self-weighting, precision of 6,000 OBRU design Self-weighting, precision of 10,000 OBRU design Not-self-weighting, precision of 6,000 OBRU design Not-self-weighting, precision of 10,000 OBRU design Not-self-weighting, precision of 6,000 total OBRU and 10,000 medicaid	102 103 98 106	1,258 2,117 1,152 1,811	5,980 9,960 5,880 9,717	3.8 6.3 3.6 5.2	4,844,013 6,931,233 4,770,353 6,758,063
OBRU designs	95	2,092	7,228	6.7	5,601,533

### References

<sup>1</sup>B. V. Shah: VMCPNLS: Program To Compute Variance Components. Research Triangle Institute. In-house report, 1979.

<sup>2</sup>G. B. Gray: Component of variance model in multi-stage stratified samples. Survey Methodology1:27-43, 1975.

<sup>3</sup>W. G. Cochran: *Sampling Techniques*. New York. John Wiley and Sons, 1977.

<sup>4</sup>National Center for Health Statistics: R. E. Folsom, Jr., R. L. Williams, and J. R. Chromy: *Optimum Design of a Medical Care Expenditure and Utilization Survey Involving a Provider Record Check.* Report No. 1725/01-06S. Research Triangle Park, N.C. Research Triangle Institute, 1980.

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# Appendix Description of cost modeling process

This appendix describes the steps required to perform the various cost modeling steps completed for the alternative designs. Examples provided in tables I-XII for this discussion are for the survey sampling operations task.

#### Table

#### Description of activity

- I Step 1. Research Triangle Institute monthly cost experience for each of the direct cost budget categories was abstracted from accounting records during the life of the project. The project activity spanned the period October 1979 through the fall of 1981.
- II Step 2. Using the monthly breakdown of project spending, monthly costs were collapsed to correspond to presurvey setup activity, rounds 1-5, and postsurvey wrapup activity periods of time.
- III Step 3. Professional staff, providing the 1980 National Medical Care Utilization and Expenditure Survey project with fiscal leadership, reviewed the round-byround cost experience to determine the level of expenditures to be associated with fixed and variable cost units of primary sampling units, segments, and reporting units (RU's). Table III shows the percents used to distribute the costs over the fixed and variable categories.
- IV Step 4. Once percent allocations were determined, these percents were applied to the actual dollars expended for each of the budget cost categories. Table IV shows actual dollar allocations for the fixed and variable modeling categories.

- V Step 5. Using various combinations of numbers expected for completed RU's, numbers of primary sampling units, and numbers of segments, the estimated costs of alternative designs were generated. Table V presents the estimated direct costs to have had only Research Triangle Institute conduct the 1980 National Medical Care Utilization and Expenditure Survey design.
- VI Procedure designed in step 5 was repeated for the 6,000-OBRU design.
- VII Procedure described in step 5 was repeated for the 10,000-OBRU design.
- VIII Step 6. In preparation for modeling the linked household unweighted design, staff reviewed the fixed and variable percent allocations used in the modeling to determine whether any refinements were to be made based on operational differences of the designs. The allocation rates for fixed and variable cost components were generated. Presented in table VIII are the dollar allocations for the fixed and variable cost categories.
- IX Step 7. Using the information prepared during step 6, staff generated the estimated costs to perform activities for the linked household unweighted design A.
- X Procedure described in step 7 was repeated for design B.
- XI Procedure described in step 7 was repeated for design C.
- XII Procedure described in step 7 was repeated for design D.

# Table I. Summary of Research Triangle Institute (RTI) cost experience for survey sampling for the National Medical Care Utilization and Expenditure Survey Household Survey, by month

			1979		19	80	
	Cost category	Total	Prior 3 months	Jan.	Feb.	Mar.	Apr.
01	Total						
	Direct technical labor						
02	On-site	\$41,138	\$7.086	\$637	\$663	\$245	\$304
03	Off-site	-	-	-	-	-	-
	Other direct cost						
04	Total	3,152	387	55	77	18	81
05	Materials and supplies	98	44	-	46	-	-
06	Services	228	48	55	7	-	-
07	Shipping and communications	213	-	-	24	18	81
80	On-site	690	-	-	-	-	-
09	Off-site	-	-	-	-	-	-
10	Consultants	-	-	-	-	-	-
11	Computer services	-	-	-	-	-	-
12	Reports and reproductions	-	-	-	-	-	-
13	Interviewer services	-	-	-	•	-	-
14	Interviewer expenses	-	-	-	-	-	-
15	Respondent incentives	-	-	-	-	-	-
16	Clerical labor	301	295	-	•	-	-
17	Clerical labor surcharge	2	-	-	•	-	-
18	Miscellaneous	-	-	-	-	-	-
19	Overtime expenses	1,620	-	-	-	-	-

NOTES: National Household Survey portion = 1.00; RTI portion = 1.00. Number of primary sampling units = 59; number of segments = 404.

	Round	Round	Round	Round	Round
ltem	1	2	3	4	5
Completed personal interviews	3,322	3,293	558	279	3,306
Completed telephone interviews	•	-	2,722	2,047	

	1980—Con.							1981						
May	June	July	Aug.	Sept.	Oct.	Nov.	Dec.	Jan.	Feb.	Mar.	Apr.	May	Other months	
•••	•••		•••								• • •			01
-	\$666	\$130 -	\$290 -	\$1,645 -	\$1,358 -	\$2,158 -	\$3,055 -	\$239 -	\$564 -	\$1,404 -	\$621 -	\$3,999 -	\$16,074 -	02 03
\$50	10	168	15	-	6	-	58	168	-	8	311	31	1,709	04
-	•	-	-	-	-	-	-	-	-	8	-	-	-	05
-		-	-	-	6	-	18	30	-	-	-	-	64	06
50	10	-	15	-	-	-	-	-	-	-	-	-	15	07
-	•	168	-	-	-	-	40	138	-	-	311	31	2	08
-	•	-	-	-	-	-	-	-	-	-	-	-	-	09
-	-	-	•	•	-	-	-	-	-	•	-	-	-	10
-	•	-	-	-	-	-	-	-	-	-	-	-	-	11
-		-	-	-	-	-	-	-	-	-	-	-	-	12
-		-		-	-	-		-	•	-	-	-	-	13
	•	-	-	-	-	-		-			-	-	-	14 15
-	•	-	-	-	-	-	-	-	-	-	-	-	6	16
•	•	-	-	-	-	-	-	-	-	-	-	-	2	17
•	-	-	-	-	-	-	-	-	•	-	-	-	-	18
-	-	-	-	-	-	-	-	-	•	-	-	-	1,620	19

 Table I.
 Summary of Research Triangle Institute (RTI) cost experience for survey sampling for the National Medical Care Utilization and Expenditure Survey Household Survey, by month—Con.

Table II. Summary of Research Triangle Institute (RTI) cost experience for survey sampling for the National Medical Care Utilization and Expenditure Survey Household Survey, rounds 1–5

Cost category	Total	Setup activity	Round 1	Round 2	Round 3	Round 4	Round 5	Wrapup activity
Total								
Direct technical labor								
On-site Off-site	\$41,138 -	\$6,732 -	\$1,899 -	\$970 -	\$3,423	\$5,213	\$6,827 -	\$16,074 -
Other direct costs								
Total	3,152	368	169	141	189	58	518	1,709
Materials and supplies	98	42	48	-	-	-	8	-
Services	228	46	64	-	6	18	30	64
Shipping and communications	213	0	42	141	15	0	. 0	15
On-site	690	-	-	-	168	40	480	2
Off-site	-	-	-	-	-	-	-	-
Consultants	-	-	-	-	-	-	-	-
Computer services	-	-	-	-	-	-	-	-
Reports and reproductions	-	-	-	-	-	-	-	-
Interviewer services	、 <b>-</b>	-	-	-	-	-	-	-
Interviewer expenses	-	-	-	-	-	-	-	-
Respondent incentives	-	-	-	-	-	-	-	-
Clerical labor	301	280	15	-	-	-	-	6
Clencal labor surcharge	2	-	-	-	-	-	-	2
Miscellaneous	-	-	-	-	-	-	-	-
Overtime expenses	1,620	-	-	-	-	-	-	1,620

NOTE: See note to table I.

Table III.	Summary of Research Triangle Institute cost experience in percent for survey sampling for the National Medical Care Utilization
and Exper	diture Survey Household Survey, rounds 1–5

			Ro	ound 1		Round 2				
	Cost category	Fixed cost	PSU cost	Segment cost	RU cost	Fixed cost	PSU cost	Segment cost	RU cost	
				,,,,,,,,,,,,,,,,,,,,,,,,,,,,,,,,,,,,,,,	Perc	ent <sup>1</sup>				
01	Total			•••		•••				
	Direct technical labor									
02	On-site	40		30	30	50			50	
03	Off-site	•••	•••	•••	•••		•••			
	Other direct costs									
04	Total	•••	•••	•••					• • •	
05	Materials and supplies	40		30	30	•••				
06	Services	40	• • •	30	30	• • •				
07	Shipping and communications	40	•••	30	30	50	•••		50	
08	On-site			•••						
09	Off-site									
10	Consultants			•••						
11	Computer services									
12	Reports and reproductions		• • •	• • •	• • •		• • •			
13	Interviewer services				• • •					
14	Interviewer expenses									
15	Respondent incentives		• • •							
16	Clerical labor	40		30	30					
17	Clerical labor surcharge			• • •	•••					
18	Miscellaneous			• • •						
19	Overtime expenses									

<sup>1</sup>Percents used to allocate fixed and per unit variable costs.

NOTE: PSU = primary sampling unit; RU = reporting unit.

 Table III.
 Summary of Research Triangle Institute cost experience in percent for survey sampling for the National Medical Care Utilization

 and Expenditure Survey Household Survey, rounds 1–5—Con.

	und 5	Ro			und 4	Ro			und 3	Ro	
RU cost	Segment cost	PSU cost	Fixed cost	RU cost	Segment cost	PSU cost	Fixed cost	RU cost	Segment cost	PSU cost	ixed cost
 					Percent <sup>1</sup>	-					
•••		•••			•••		•••			•••	•••
25			75	15			85	25	•••		75
•••		•••		•••		•••	•••	•••		•••	•••
25			75								
25			75	15			85	25		• • •	75
25			75		•••			25		•••	75
25			75	15	•••		85	25			75
					• • •						
								• • •			
					•••			• • •	•••		
				• • •	• • •			•••			
• • •	•••				• • •		•••	•••		• • •	
			• • •		•••		•••	• • •	• • •		•••
25		• • •	75	• • •	• • •	•••		• • •			
25			75	• • •	•••	• • •	•••	• • •		• • •	• • •
•••	• • •		•••	•••	• • •	•••	•••	• • • •	•••		• • •
• • •							• • •	• • •	• • •		

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 Table IV.
 Summary of Research Triangle Institute cost experience for survey sampling for the National Medical Care Utilization and Expenditure

 Survey Household Survey, rounds 1–5, by type of cost

			Roui	nd 1		Round 2				
	Cost category	Fixed cost	PSU cost	Segment cost	RU cost	Fixed cost	PSU cost	Segment cost	RU cost	
01	Total	\$3,667.20	-	\$6.81	\$0.83	\$555.50	-	-	\$0.17	
	Direct technical labor									
02 03	On-site Off-site	3,452.40	-	6.41 -	0.78	485.00	-	-	0.15	
	Other direct costs									
04	Total	214.80	-	0.40	0.50	70.50	-	-	0.02	
05	Materials and supplies	36.00	-	0.07	0.01	-	-	-	-	
06	Services	44.00	-	0.08	0.01	-	-	-	-	
07	Shipping and communications	16.80	-	0.03	-	70.50	-	-	0.02	
80	On-site	-	-	-	-	-	-	-	-	
09	Off-site	-	-	-	-	-	-	-	-	
10	Consultants	-	-	-	•	-	-	-	-	
11	Computer services	-	-	-	-	-	-	-	-	
12	Reports and reproductions	-	-	-	-	-	-	-	-	
13	Interviewer services	-	-	-	-	-	-	-	-	
14	Interviewer expenses	-	-	-	-	-	-	-	-	
15	Respondent incentives	-	-	-	-	-	-	-	-	
16	Clerical labor	118.00	-	0.22	0.03	-	-	-	-	
17	Clerical labor surcharge	-	-	•	-	-	-	-	-	
18	Miscellaneous	-	-	-	-	-	-	-	-	
19	Overtime expenses	-	-	-	-	-	-	-	-	

,

NOTES: PSU = primary sampling unit; RU = reporting unit. Number of primary sampling units = 59; number of segments = 404.

ltem	Round 1	Round 2	Round 3	Round 4	Round 5
Completed personal	3.322	3.293	558	279	3,306
Completed telephone	0,522				0,000
interviews	-	-	2,722	2,047	
Cost			Percent		
Total	100	100	100	100	100
Fixed	40	50	75	85	75
PSU	30		-	-	-
Case	30	50	25	15	25

Table IV.	Summary of Research Triangle Institute cost experience for survey sampling for the National Medical Care Utilization and Expenditure
Survey H	ousehold Survey, rounds 1–5, by type of cost—Con.

		d 5	Roun			nd 4			Round 3					
	RU cost	Segment cost	PSU cost	Fixed cost	RU cost	Segment cost	PSU cost	Fixed cost	RU cost	Segment cost	PSU cost	Fixed cost		
01	\$1.90	-	-	\$18,846.00	\$0.34	-	-	\$4,480.35	\$0.28	•	-	\$2,709.00		
02 03	1.73 -	:	-	17,175.75	0.34	-	-	4,431.05	0.26	:	-	2,567.25		
04	0.17		-	1,670.25	-	-	-	49.30	0.01		-	141.75		
05	-	-	-	6.00	-	-	-	-	-	-	-	-		
06	0.01	-	-	70.50	-	-	-	15.30	-	-	-	4.50		
07	-	-	-	11.25	-	-	•	-	-	-	•	11.25		
08	0.04	-	-	361.50	-	-	-	34.00	0.01	-	-	126.00		
09	-	-	-		-	-	-	-	-	-	-	-		
10	-	-	-	-	-	-	-	-	-	-	-	-		
11	-	-	-	-	-	-	-	-	-	-	-	-		
12	-	-	-	-	-	-	-	-	-	-	-	-		
13	-	-	-	-	-	-	-	-	-	-	•	•		
14	-	-	-	-	-	-	-	-	-	-	-	-		
15	-	-	-	-	-	-	-	-	-	-	-	-		
16	-	-		4.50	-	-	-	-	-	-	-	•		
17	-	-	-	1.50	-	-	-	-	-	-	-	-		
18	-	-	-	-	-	-	-	-	-	-	-	-		
19	0.12	-	-	1,215.00	-	-	-	-	-	-	-	-		

### Table V. Summary of estimated costs for survey sampling with the Research Triangle Institute design component of the 1980 NMCUES

				Round 1						Round 2				
	Cost category	Total	Total	Fíxed cost	PSU cost	Seg- ment cost	RU cost	Total	Fixed cost	PSU cost	Seg- ment cost	RU cost		
01	Total	\$58,147	\$14,642	\$3,667	-	\$5,508	\$5,467	\$1,656	\$556	-	-	\$1,100		
	Direct technical labor													
02	On-site	54,001	13,784	3,452	-	5,185	5,147	1,445	485	-	-	960		
03	Off-site	-	-	-	-	-	-	-	•	-	-	-		
	Other direct costs													
04	Total	4,154	858	215	-	323	320	219	71	-	-	140		
05	Materials and supplies	154	144	36	-	54	54	-	-	-	-	-		
06	Services	322	176	44	-	66	66	-	-	-	-	-		
07	Shipping and communications	322	67	17	-	25	25	219	71	-	-	140		
08	On-site	857	-	-	-	-	-	-	-	-	-	-		
09	Off-site	-	-	-	-		-	-	-	-	-	-		
10	Consultants	-	-	-	-	-	-	-	-	-	-	-		
11	Computer services	-	-	-	-		-	-	-	-	-	-		
12	Reports and reproductions	-	-	-	-	-	-	-	-	-	-	-		
13	Interviewer services	-	-	-	-	•	-	-	-	-	-	-		
14	Interviewer expenses	-	-	-	-	•	-	-	-	-	-	-		
15	Respondent incentives	-	-	-	-	•	-	-	-	-	-	-		
16	Clerical labor	479	471	118	-	177	176	•	-	-	-	-		
17	Clerical labor surcharge	3	-	-	-	-	-	-	-	-	-	•		
18	Miscellaneous	-	-	•	-	•	-	-	-	-	-	-		
19	Overtime expenses	2,019	-	-	-	-	-	•	-	-	-	-		

NOTES: Number of primary sampling units (PSU's) = 108; number of segments = 809; RU = reporting unit. Data are based on NMCUES fixed and per unit allocations.

ltem	Round 1	Round 2	Round 3	Round 4	Round 5
Completed personal Interviews	6,603	6,519	1,110	547	6,561
Completed telephone Interviews	-	-	5,418	4,012	-

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Table V.	Summary of estimated costs for survey	ey sampling with the Research Triangle	Institute design component of the 1980 NMCUES-Con.
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	Round 3					Round 4					Round 5				
Total	Fixed cost	PSU cost	Seg- ment cost	RU cost	Total	Fixed cost	PSU cost	Seg- ment cost	RU cost	Total	Fixed cost	PSU cost	Seg- ment cost	RU cost	
\$4,506	\$2,709	-	-	\$1,797	\$6,030	\$4,480	-	-	\$1,550	\$31,313	\$18,846	-		\$12,467	01
4,270	2,567 -	:	-	1,703	5,964 -	4,431 -	-	-	1,533 -	28,538 -	17,176 -	- -	-	11,362 -	02 03
236	142	-	-	94	66	49	-	-	17	2,775	1,670	-		1,105	04
-	•	-	-	-	-	-	-	-	-	10	6	-	-	4	05
8	5	-	-	3	20	15	-	-	5	118	71	-	-	47	06
18	11	-	-	7	-	-	-	-	-	18	11	•	-	7	07
210	126	-	-	84	46	34	-	-	12	601	362	-	-	239	08
-	-	-	-	-	-	-	-	-	-	-	-	-	-	-	09
-	-	-	-	-	-	-	-	-	-	-	-	-	-	-	10
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-		-	-	-	-	-	-	-	-	2,019	1,215	-	-	804	19

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#### Table VI. Summary of estimated costs for survey sampling with the 6,000-respondent originating base reporting unit unlinked design

				Round 2								
	Cost category	Total	Total	Fixed cost	PSU cost	Seg- ment cost	RU cost	Total	Fixed cost	PSU cost	Seg- ment cost	RU cost
01	Total	\$56,782	\$14,005	\$3,667	-	\$5,106	\$5,232	\$1,610	\$556	-	-	\$1,054
	Direct technical labor											
02 03	On-site Off-site	52,732 -	13,184 -	3,452 -	-	4,807 -	4,925 -	1,405 -	485	-	-	920 -
	Other direct costs											
04	Total	4,049	820	215	-	299	306	205	71	-	-	134
05	Materials and supplies	147	137	36	-	50	51	-	-	-	-	-
06	Services	312	168	44	-	61	63	-		-	-	-
07	Shipping and communications	305	64	17	-	23	24	205	71	-	-	134
08	On-site	842	-	-	-	-	-	-	-	-	-	-
09	Off-site	-	-	-	-	-	-	-	-	-	-	-
10	Consultants	-	-	-	-	-	-	-	-	-	-	-
11	Computer services	-	-	-	-	•	•	-	-	-	-	-
12	Reports and reproductions	-	-	-	-	-	-	-	-	-	-	-
13	Interviewer services	-	-	-	-	-	-	-	•	-	-	-
14	Interviewer expenses	-	-	-	-	-	-	-	-	-	•	•
15	Respondent incentives	-	-	-	-	-	-	-	-	-	-	-
16	Clerical labor	458	450	118	-	164	168	-	-	•	•	-
17	Clerical labor surcharge	3	-	-	-	-	-	•	-	•	-	-
18	Miscellaneous	-	-	-	-	-	•	-	-	-	-	-
19	Overtime expenses	1,984	-	-	-	-	-	•	-	-	-	-

NOTES: Number of primary sampling units (PSU's) = 102; number of segments = 750; RU = reporting unit. Data are based on NMCUES fixed and per unit allocations.

ltem	Round 1	Round 2	Round 3	Round 4	Round 5
Completed personal interviews	6,319	6,247	1,062	524	6,278
Completed telephone interviews	-	-	5,185	3,839	-

Table VI.	Summary	of estimated cost	s for survev san	plind	with the 6.000-res	pondent originatin	g base reportin	g unit unlinked designCon.

Round 5						R		Round 3							
RU ost	nt	Seg mer cos	PSU cost	Fixed cost	Total	RU cost	Seg- ment cost	PSU cost	Fixed cost	Total	RU cost	Seg- ment cost	PSU cost	Fixed cost	Total
,929		-	-	\$18,846	\$30,775	\$1,483	-	-	\$4,480	\$5,963	\$1,720		-	\$2,709	\$4,429
),872 -		-	-	17,176 -	28,048 -	1,467	-		4,431	5,898 -	1,630 -	-	-	2,567	4,197 -
,057		-	-	1,670	2,727	16		-	49	65	90	-		142	232
4		-	-	6	10	-	-	-	-	-	-	-	-	-	-
45		-	-	71	116	5	-	-	15	20	3	-	-	5	8
7		-	•	11	18	-	-	•	-	-	7	-	•	11	18
229		-		362	591	11	-	•	34	45	80	-	-	126	206
-		-	-	-	-	-	-	-	-	-	-	-	-	-	-
-		-	-	-	-	-	•	-	-	-	-	-	-	-	-
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769		-	-	1,215	1,984	-	-	-	-	•	-	-	-	-	-

#### Table VII. Summary of estimated costs for survey sampling with the 10,000-respondent originating base reporting unit unlinked design

				96 \$3,667 - 71 3,452 -  24 215 - 26 36 - 51 44 -			Round 2					
	Cost category	Total	Total			Seg- ment cost	RU cost	Total	Fixed cost	PSU cost	Seg- ment cost	RU cost
01	Total	\$74,462	\$20,896	\$3,667	-	\$8,510	\$8,719	\$2,310	\$556	•	-	\$1,754
	Direct technical labor											
02	On-site	69,142	19,671	3.452	-	8,011	8,208	2.016	485	-	-	1,531
03	Off-site	-	-	-	-	-	•	-	•	-	•	-
	Other direct costs											
04	Total	5,318	1,224	215	-	498	511	294	71	-	-	223
05	Materials and supplies	218	206	36	-	84	86	-	-	-	-	
06	Services	429	251	44	-	102	105	-	-	-	-	
07	Shipping and communications	436	96	17	-	39	40	294	71	-	-	223
08	On-site	1,055	-	-	-	-	-	-	-	-	-	
09	Off-site	-	-	-	-	-	-	-	-	-	-	
10	Consultants	-	-	-	-	-	-	-	-	-	-	-
11	Computer services	-	-	-	-	-	-	-	-	-	-	-
12	Reports and reproductions	-	-	-	-	-	-	-	-	-	-	-
13	Interviewer services	-	-	-	-	-	-	-	-	-	-	-
14	Interviewer expenses	-	-	-	-	-	-	-	-	-	-	-
15	Respondent incentives	-	-	-	-	-	-	-	-	-	-	-
16	Clerical labor	683	673	118	-	274	281	•	-	-	-	-
17	Clerical labor surcharge	4	-	-	-	-	-	-	-	-	-	-
18	Miscellaneous	-	-	-	-	-	-	-	-	-	-	-
19	Overtime expenses	2,497	-	-	-	-	-	-	-	-	-	-

NOTES: Number of primary sampling units (PSU's) = 102; number of segments = 1,250; RU = reporting unit. Data are based on NMCUES fixed and per unit allocations.

ltem	Round 1	Round 2	Round 3	Round 4	Round 5
Completed personal interviews	10,531	10,397	1,770	873	10,464
Completed telephone interviews		-	8,641	6,398	-

Table VII.	Summary of estimated costs for survey sampling with the	e 10,000-respondent originating base reporting unit unlinked design-Con.

	Round 5				Round 4					Round 3					
	RU cost	Seg- ment cost	PSU cost	Fixed cost	Total	RU cost	Seg- ment cost	PSU cost	Fixed cost	Total	RU cost	Seg- ment cost	PSU cost	Fixed cost	Total
3 01	\$19,883	-	-	\$18,846	\$38,729	\$2,472	•	•	\$4,480	\$6,952	\$2,866		-	\$2,709	\$5,575
	18,121	-	-	17,176	35,297	2,444	-	-	4,431	6,875	2,716	-	-	2,567	5,283
- 03	-	-	-	-	-	-	-	-	-	•	-	-	•	-	-
2 04	1,762	-	-	1,670	3,432	27		-	49	76	150	-	-	142	292
3 05	6	-	-	6	12	-	-	-	-	-	-	-	-	-	-
	74	-	-	71	145	8	-	-	15	23	5	-	-	5	10
	12	•	-	11	23	-	-	-	•	-	12	-	-	11	23
1 08	381	-	-	362	743	19	-	-	34	53	133	-	-	126	259
- 09	-	-	-	-	-	-	-	-	-	-	-	-	-	-	-
- 10	-	-	-	-	-	•	-	-	-	-	•	-	-	-	-
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2 17 - 18	2	-	-	2	4	•	•		-	-	-	-	-	•	
	1,282	-	-	- 1,215	- 2,497	-	-	-	-	-	-	-	-		

#### Table VIII. Summary of costs for survey sampling for the linked household design

			Rou	nd 1			Ro	und 2	
	Cost category	Fixed cost	PSU cost	Segment cost	RU cost	Fixed cost	PSU cost	Segment cost	RU cost
01	Total	\$3,111.20	-	\$5.78	\$0.70	\$555.50	-	-	\$0.17
	Direct technical labor								
02	On-site	2,471.80	-	4.59	0.56	485.00	-	-	0.15
03	Off-site	-	-	-	-	-	-	-	-
	Other direct costs								
04	Total	639.60	-	1.19	0.14	70.50	-	-	0.02
05	Materials and supplies	36.00	-	0.07	0.01	-	-	-	-
06	Services	44.00	-	0.08	0.01	-	-	-	-
07	Shipping and communications	16.80	-	0.03	-	70.50	-	-	0.02
08	On-site	-	-	-	-	-	-	-	-
09	Off-site	-	-	-	-	-	-	-	
10	Consultants	-	-	-	-	-	-	-	
11	Computer services	-	-	-	-	-	-	-	-
12	Reports and reproductions	-	-	-	-	-	-	-	
13	Interviewer services	-	-	-	-	-	-	-	
14	Interviewer expenses	-	-	-	-	-	-	-	-
15	Respondent incentives	-	-	-	-	-	-	-	-
16	Clerical labor	457.60	-	0.85	0.10	-	-	-	-
17	Clerical labor surcharge	85.20	-	0.16	0.02	-	-	-	-
18	Miscellaneous	-	-	-	-	-	-	-	-
19	Overtime expenses	-	-	•	-	-	-	-	-

NOTES: PSU = primary sampling unit; RU = reporting unit. Number of primary sampling units = 59; number of segments = 404.

ltem	Round 1	Round 2	Round 3	Round 4	Round 5
Completed personal interviews Completed telephone	3,322	3,293	558	279	3,306
interviews	-	-	2,722	2,047	•

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Table VIII. S	Summary of costs	for survey s	sampling for	r the linked h	nousehold de	∋sign—Con.
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	d 5	Roun			nd 4	Roui			nd 3	Roui	
RU cos	Segment cost	PSU cost	Fixed cost	RU cost	Segment cost	PSU cost	Fixed cost	RU cost	Segment cost	PSU cost	Fixed cost
\$1.9	-		\$18,846.00	\$0.34	-	-	\$4,480.35	\$0.28	-	-	62,709.00
1.7			17,175.75	0.34	-	-	4,431.05	0.26	-		2,567.25
	-	-	-	-	-	-	-	-	-	-	-
0.1	-	-	1,670.25	-	-	-	49.30	0.01	-	-	141.75
	-	-	6.00	-	-	-	-	-	-	-	-
0.0	-	-	70.50	-	-	-	15.30	-	-	-	4.50
	-	-	11.25	-	-	-		-	-	-	11.25
0.0		-	361.50	-	-	-	34.00	0.01	-	-	126.00
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	-	-	4.50	-	-	-	•	-	-	-	-
	-	-	1.50	-	-	-	-	-	-	-	•
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0.1	-	•	1,215.00	-	-	-	-	-	-	-	•

#### Table IX. Summary of estimated costs for survey sampling for the linked household design A

				ŀ	Round 1			Round 2					
	Cost category	Total	Total	Fixed cost	PSU cost	Seg- ment cost	RU cost	Total	Fixed cost	PSU cost	Seg- ment cost	RU cost	
01	Total	\$55,461	\$13,078	\$3,111	-	\$5,637	\$4,330	\$1,582	\$556	•	-	\$1,026	
	Direct technical labor												
02	On-site	49.575	10,390	2,472	-	4,478	3,440	1,381	485	-	-	896	
03	Off-site	-	-	-,	-	-	-	-	•	-	-	-	
	Other direct costs												
04	Total	5,887	2,689	640	-	1,159	890	201	71	-	-	130	
05	Materials and supplies	161	151	36	-	65	50	-	-	-	-	-	
06	Services	328	185	44	-	80	61	-	-	-	-	-	
07	Shipping and communications	307	70	17	-	30	23	201	71	-	-	130	
08	On-site	834	-	-	-	-	-	-	-	-	-	-	
09	Off-site		-	-	-	-	-	-	-	-	-	-	
10	Consultants	-	-	-	-	-	-	-	-	-	-	-	
11	Computer services	-	-	-	-	-	-	-	-	-	-	-	
12	Reports and reproductions	-	-	-	•	-	-	-	-	-	-	-	
13	Interviewer services	-	-	-	-	-	-	-	-	-	-	-	
14	Interviewer expenses	-	-	-	-	-	-	-	-	-	-	-	
15	Respondent incentives	-	-	-	-	-	-	-	-	-	-	-	
16	Clerical labor	1,932	1,924	458	-	829	637	-	-	-	-	-	
17	Clerical labor surcharge	361	358	85	-	154	119	-	-	-	-	-	
18	Miscellaneous	-	-	-	-	-	-	-	-	-	-	-	
19	Overtime expenses	1,965	-	•	-	-	-	•	-	-	-	-	

NOTES: Number of primary sampling units (PSU's) = 100; number of segments = 976; RU = reporting unit.

ltem	Round 1	Round 2	Round 3	Round 4	Round 5
Completed personal interviews Completed telephone	6,165	6,084	1,036	511	6,126
interviews	-	-	5,058	3,746	-

 Table IX.
 Summary of estimated costs for survey sampling for the linked household design A--Con.

			und 5	Ro				ound 4	R				ound 3	R	
	RU cost	Seg- ment cost	PSU cost	Fixed cost	Total	RU cost	Seg- ment cost	PSU cost	Fixed cost	Total	RU cost	Seg- ment cost	PSU cost	Fixed cost	Total
01	\$11,641	-	-	\$18,846	\$30,487	\$1,447	-	-	\$4,480	\$5,927	\$1,678	-	-	\$2,709	\$4,387
02 03	10,609 -	-		17,176 -	27,785	1,431	-	- -	4,431 -	5,862 -	1,590		-	2,567 -	4,157 -
04	1,032	-		1,670	2,702	16		-	49	65	88		-	142	230
05	4	-	-	6	10	-	-	-	-	-		-	-	-	-
06	44	-	-	71	115	5	-	-	15	20	3	-	-	5	8
07	7	-	-	11	18	-	-	•	-	-	7	-	-	11	18
08	223			362	585	11	-		34	45	78	-	-	126	204
09	-	-	-	-	-	-	-	-	-	-	•	-	-	-	-
10	-	-	-	-	-	-	-	-	-	-	-	-	-	-	-
11	•	-	-	-	-	-	-	-	•	•	•	-	•	-	-
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19	750	-	-	1,215	1,965		-	-	-			-	-		-

### Table X. Summary of estimated costs for survey sampling for the linked household design B

				F	Round 1			Round 2				
	Cost category	Total	Total	Fixed cost	PSU cost	Seg- ment cost	RU cost	Total	Fixed cost	PSU cost	Seg- ment cost	RU cost
01	Total	\$54,009	\$12,517	\$3,111	-	\$5,319	\$4,087	\$1,524	\$556	•	-	\$968
	Direct technic <del>a</del> l labor											
02	On-site	48,310	9,944	2,472	-	4,226	3,246	1,331	485	-	-	846
03	Off-site	•	-	•	-	•	-	-	-	-	-	-
	Other direct costs											
04	Total	5,701	2,574	640	-	1,094	840	194	71	-	-	123
05	Materials and supplies	154	145	36	-	62	47	-	-	-	-	
06	Services	317	177	44	-	75	58	-	-	-	-	-
07	Shipping and communications Travel:	298	68	17	-	29	22	194	71	-	-	123
08	On-site	817	-	-	-	-	-	-	-	-	-	-
09	Off-site	-	-		-	-	-	-		-	-	-
10	Consultants	-	-	-	-	-	-	-	-	-	-	-
11	Computer services	-	-	-	-	-	-	-	-	-	-	-
12	Reports and reproductions	-	-	-	-	-	-	-	-	-	-	-
13	Interviewer services	-	-	-	-	-	-	-	-	-	-	-
14	Interviewer expenses	-	-	-	-	-	-	-	-	-	-	-
15	Respondent incentives	-	-	-	-	-	-	-	-	-	-	-
16	Clerical labor	1,849	1,841	458	-	782	601	۰.	-	-	-	-
17	Clerical labor surcharge	346	343	85	-	146	112	2	-	-	-	-
18	Miscellaneous	-	-	•	-	-	-	-	•	-	-	-
19	Overtime expenses	1,923	-	-	-	-	-	•	-	-	-	-

NOTES: Number of primary sampling units (PSU's) = 200; number of segments = 921; RU = reporting unit.

ltem	Round 1	Round 2	Round 3	Round 4	Round 5
Completed personal interviews Completed telephone	5,818	5,741	978	482	5,781
interviews	-	-	4,773	3,535	

Table X. St	ummary of estimated	costs for survey sa	ampling for the li	inked household desig	an BCon.
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	F	lound 3				F	Round 4				Ro	und 5			
Total	Fixed cost	PSU cost	Seg- ment cost	RU cost	Total	Fixed cost	PSU cost	Seg- ment cost	RU cost	Total	Fixed cost	PSU cost	Seg- ment cost	RU cost	
\$4,292	\$2,709	-	-	\$1,583	\$5,845	\$4,480	-	-	\$1,365	\$29,831	\$18,846	-	-	\$10,985	01
4,067 -	2,567 -	-		1,500 -	5,781 -	4,431	-	-	1,350 -	27,187 -	17,176	- -	-	10,011 -	02 03
225	142	-	•	83	64	49	-	-	15	2,644	1,670	-	-	974	04
-	-	-	-	-	-	-	-	-	-	9	6	-	-	3	05
8 18	5 11	-	-	3 7	20	15	-	-	5	112 18	71 11	-	-	41 7	06 07
200	126	-	-	74	44	34	-	-	10	573	362	-	-	211	08
-	-	-	-	-	-	•	-	-	-	-	-	-	-	-	09
-	-	-	-	-	-	-	-	-	-	-	-	-	-	-	10
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-	•	-	-	-	-	-	-	-	-	1,923	1,215	-	-	708	19

#### Table XI. Summary of estimated costs for survey sampling for the linked household design C

				ŀ	Round 1					Round 2	2	
	Cost category	Total	Total	Fixed cost	PSU cost	Seg- ment cost	RU cost	Total	Fixed cost	PSU cost	Seg- ment cost	RU cost
01	Total	\$72,690	\$19,742	\$3,111	-	\$9,909	\$7,228	\$2,269	\$556	-	-	\$1,713
	Direct technical labor											
02	On-site	64.593	15,688	2,472	-	7,474	5,742	1,981	485	-	-	1,496
03	Off-site	-	-	•	-	-	-	•	•	-	-	-
	Other direct costs											
04	Total	8,105	4,060	640	-	1,934	1,486	288	71	-	-	217
05	Materials and supplies	241	229	-	-	109	84	-	-	-	-	-
06	Services	456	279	-	-	133	102	-	-	-	-	
07	Shipping and communications Travel:	441	107	71	-	51	39	288	71	-	•	217
08	On-site	1,043	-	-	-	-	-	-	-	-	-	-
09	Off-site	-	-	-	-	-	-	-	-	-	-	-
10	Consultants	-	-	-	-	-	-	-	-	-	-	-
11	Computer services	-	-	-	-	-	-	-	-	-	-	-
12	Reports and reproductions	-	-	-	-	-	-	-	-	-	-	-
13	Interviewer services	-	-	-	-	-	-	-	-	-	-	-
14	Interviewer expenses	-	-	-	-	-	-	-	-	-	-	-
15	Respondent incentives	-	-	-	-	-	-	-	-	-	-	-
16	Clerical labor	2,915	2,905	458	-	1,384	1,063	-	-	-	-	-
17	Clerical labor surcharge	545	541	85	-	258	198	-	-	-	-	-
18	Miscellaneous	-	-	-	-	-	-	-	-	-	-	-
19	Overtime expenses	2,468	-	•	-	-	-	-	-	-	-	-

NOTES: Number of primary sampling units (PSU's) = 100; number of segments = 1,629; RU = reporting unit.

ltem	Round 1	Round 2	Round 3	Round 4	Round 5
Completed personal interviews.	10,290	10,154	1,729	853	10,225
Completed telephone interviews		-	8,443	6,252	-

Table XI. Su	mary of estimated costs fo	r survey sampling for the	Iinked household design C—Con.
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			Round 4 Round 5								ound 3	R			
	RU cost	Seg- ment cost	PSU cost	Fixed cost	Total	RU cost	Seg- ment cost	PSU cost	Fixed cost	Total	RU cost	Seg- ment cost	PSU cost	Fixed cost	Total
01	\$19,429	-	-	\$18,846	\$38,275	\$2,415	-	-	\$4,480	\$6,895	\$2,800	-	-	\$2,709	\$5,509
02 03	17,707 -	-	-	17,176	34,883 -	2,389 -	-		4,431 -	6,820	2,654 -	-	-	2,567 -	5,221 -
04	1,722	-	-	1,670	3,392	27	-	-	49	76	147	-	-	142	289
05	6	-	-	6	12	-	-	-	-	-	-	-	-	-	-
06	73	-	-	71	144	8	-	-	15	23	5	-	-	5	10
07	12	-	-	11	23	-	-	-	-	-	12	-	-	11	23
08	373	-	-	362	735	18	-	-	34	52	130	-	-	126	256
09	-	-	-	-	-	-	-	-	-	•	-	-	-	-	-
10	-	-	-	-	-	-	-	-	-	-	-	-	-	-	-
11	-	-	-	-	-	-	-	-	-	•	-	-	-	-	-
12	-	-	-	-	-	•	-	-	-	•	-	-	-	-	-
13	-	-	-	-	-	•	-	-	-	-	-	-	-	-	-
14	-	-	-	-	-	-	-	-	-	-	-	-	-	-	-
15	-	-	-	-	-	-	-	-	-	-	-	-	-	-	-
16 17	5 2	-	-	5 2	10 4	-	-	-	-	-	-	-	-	-	-
18	2	-	-	2	4	-	-	-	-	-	-	-	-	-	-
19	- 1,253	-	-	1,215	2,468	-	-	-	-			-	-	-	-

# Table XII. Summary of estimated costs for survey sampling for the linked household design D

			F	Round 1		Round 2						
	Cost category	Total	Total	Fixed cost	PSU cost	Seg- ment cost	RU cost	Total	Fixed cost	PSU cost	Seg- ment cost	RU cost
01	Total	\$69,002	\$18,318	\$3,111	-	\$8,600	\$6,607	\$2,122	\$556	-	-	\$1,566
	Direct technical labor											
02 03	On-site Off-site	61,373 -	14,553 -	2,472 -	-	6,832 -	5,249 -	1,852 -	485 -	-	-	1,367 -
	Other direct costs											
04	Total	7,632	3,766	640	-	1,768	1,358	270	71	-	•	199
05 06	Materials and supplies	226 428	214 259	36 44	-	100 122	76 93	-	-	-	-	-
07	Shipping and communications Travel:	413	99	17	-	46	36	270	71	-	-	199
08	On-site	999	-	-	-	-	-	-	-	-	-	-
09	Off-site	-	-	-	-	-	-	-	-	-	-	•
10	Consultants	-	-	-	•	-	-	-	-	-	-	-
11	Computer services	-	-	-	-	-	-	-	-	-	-	-
12	Reports and reproductions	-	-	-	-	-	-	-	-	-	-	-
13	Interviewer services	-	-	-	-	-	-	-	-	-	•	-
14	Interviewer expenses	-	-	-	-	-	-	-	-	-	-	-
15	Respondent incentives	-	-	-	-	-	-	-	-	-	-	-
16	Clerical labor	2,704	2,695	458	-	1,265	972	-	-	-	-	-
17	Clerical labor surcharge	505	502	85	-	236	181	-	-	•	-	-
18	Miscellaneous	-	-	-	-	-	-	-	-	-	-	-
19	Overtime expenses	2,360	-	-	-	-	-	-	-	-	-	-

NOTES: Number of primary sampling units (PSU's) = 200; number of segments = 1,489; RU = reporting unit.

ltem	Round 1	Round 2	Round 3	Round 4	Round 5
Completed personal interviews Completed telephone	9,406	9,282	1,580	779	9,346
interviews	-	-	7,717	5,716	-

	F	Round 3				F	Round 4			Round 5						
Total	Fixed cost	PSU cost	Seg- ment cost	RU cost	Total	Fixed cost	PSU cost	Seg- ment cost	RU cost	Total	Fixed cost	PSU cost	Seg- ment cost	RU cost		
\$5,269	\$2,709	-	-	\$2,560	\$6,688	\$4,480	-	•	\$2,208	\$36,605	\$18,846	-	-	\$17,759	01	
4,993 -	2,567 -		-	2,426	6,614 -	4,431 -	-	•	2,183	33,361	17,176 -		- -	16,185	02 03	
276	142	-	-	134	76	49	-	-	24	3,244	1,670	•	-	1,574	04	
-	-	-	-	-	-	-	-	-	-	12	6	-	-	6	05	
9	5	-	-	4	23	15	-	-	8	137	71	-	-	66	06	
22	11	-	-	11	-	-	-	-	-	22	11	-	-	11	07	
245	126	-	-	119	51	34	-	-	17	703	362	-	-	341	08	
•	-	-	-	-	-	-	-	-	-	-	-	-	-	-	09	
•	-	-	-	-	-	-	-	-	-	-	-	-	-	-	10	
•	•	-	-	-	-	•	-	-	-	-	-	-	-	-	11	
-	-	•	-	-	-	-	-	-	-	-	-	-	-	-	12	
•	-	•	-	-	-	-	-	-	•	-	-	-	-	-	13	
-	-		-	-	-	•	-	-		-	-	-	-	-	14 15	
-		-	-				-	-		-	5	-	-	4	16	
-	-	-	-	-	-	-	-	-	-	-	2	-	-	1	17	
-	-	•	-	-	-	-	-	-	-	-	-	-	-	-	18	
-	-	•	-	-	-	-	-	-	-	2,360	1,215	-	•	1,145	19	

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